

WEAK CONVERGENCE TO BROWNIAN EXCURSION

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ABSTRACT. We show that Brownian motion conditioned to be positive and tied-down is Brownian excursion. For a random walk defined by weakly dependent random variables, a conditional limit theorem is proved under some conditions on the path.

1. Introduction. Our purpose in this paper is to investigate relationship between Brownian motion and Brownian excursion.

Let $\{W(t): t \geq 0\}$ be a standard Brownian motion. Let

$$n_s(x) = (2\pi s)^{-1/2} \exp(-x^2/2s), \quad N_s(a, b) = \int_a^b n_s(x) dx, \quad \text{and}$$

$$g(t, x, y) = n_t(y-x) - n_t(y+x).$$

Brownian excursion, $W_0^+ = \{W_0^+(t): 0 \leq t \leq 1\}$ is a continuous, nonhomogeneous Markov process. The transition density is given by

$$P[W_0^+(t) \in dy] = p_0^+(0, 0, t, y) = \frac{2y^2 \exp(-y^2/2t(1-t))}{(2\pi t^3(1-t)^3)^{1/2}} dy$$

for $0 < t \leq 1$; for $0 < s < t < 1$ and $x, y > 0$,

$$P[W_0^+(t) \in dy \mid W_0^+(s) = x] = p_0^+(s, x, t, y) dy$$

$$= g(t-s, x, y) \{(1-s)/(1-t)\} \frac{y \exp(-y^2/2(1-t))}{x \exp(-x^2/2(1-s))} dy.$$

Let $\{W^+(t): 0 \leq t \leq 1\}$ be Brownian meander.

According to Durrett-Iglehart-Miller [2] we introduce some notation, for $0 \leq t \leq 1$, let $m(t) = \inf\{W(s); 0 \leq s \leq t\}$ and $M(t) = \sup\{W(s); 0 \leq s \leq t\}$. Furthermore let $C = C[0, 1]$ be the space of continuous functions on $[0, 1]$ and \mathcal{C} be the topological σ -field of C induced by the supremum metric $d(\cdot, \cdot)$. For a random function Y in (C, \mathcal{C}) let Q be the probability measure of Y on (C, \mathcal{C}) and Q_A the probability measure restricted by a Borel set $A \in \mathcal{C}$ with $Q(A) > 0$ which is defined by $Q_A(A) = Q(A)/Q(A)$ for $A \in \mathcal{C}$. Then we can define a random function $Y|A$ with the probability measure Q_A as the restriction of Y to $Y^{-1}(A)$. We now define the following conditioned random function of (C, \mathcal{C}) , for any $\epsilon > 0$,

$$\bar{W}_\varepsilon = W | \{m(1) > -\varepsilon, W(1) < \varepsilon\}.$$

This process is Markov by virtue of Lemma (1.5) in [2].

2. Convergence of conditioned Brownian motion to Brownian excursion.

In this section we prove the following.

Theorem (2.1). $\bar{W}_\varepsilon \Rightarrow W_0^+$ as $\varepsilon \rightarrow 0$.

Proof. We first show that the finite dimensional distributions converge. Since \bar{W}_ε and W_0^+ are Markov, it suffices to show that the probability transition densities converge. For $0 < t < 1$ and $y > 0$, we have

$$\begin{aligned} \bar{p}_\varepsilon(0, 0, t, y) &= P[W(t) \in dy | m(1) > -\varepsilon, W(1) < \varepsilon] \\ &= \frac{P[W(t) \in dy, m(t) > -\varepsilon] P^y[m(1-t) > -\varepsilon, W(1-t) < \varepsilon]}{P[m(1) > -\varepsilon, W(1) < \varepsilon]} \\ &= \frac{g(t, \varepsilon, y + \varepsilon) \int_{-\varepsilon}^{\varepsilon} g(1-t, y + \varepsilon, u + \varepsilon) du dy}{\int_{-\varepsilon}^{\varepsilon} g(1, \varepsilon, u + \varepsilon) du}. \end{aligned}$$

Using L'Hospital's rule twice on the ratio above gives

$$\lim_{\varepsilon \rightarrow 0} \bar{p}_\varepsilon(0, 0, t, y) = \frac{2y^2 \exp(-y^2/2t(1-t))}{(2\pi t^3(1-t)^3)^{1/2}} = \bar{p}_0(0, 0, t, y).$$

For $0 < s < t \leq 1$ and $x, y > 0$

$$\begin{aligned} \bar{p}_\varepsilon(s, x, t, y) dy &= P[\bar{W}_\varepsilon(t) \in dy | \bar{W}_\varepsilon(s) = x] \\ &= \{P[W(s) \in dx, m(s) > -\varepsilon] P^x[m(t-s) > -\varepsilon, W(t-s) \in dy] \\ &\quad \times P^y[m(1-t) > -\varepsilon, W(1-t) < \varepsilon]\} / \{P[W(s) \in dx, m(s) > -\varepsilon] \\ &\quad \times P^x[m(1-s) > -\varepsilon, W(1-s) < \varepsilon]\} \\ &= \frac{g(t-s, x + \varepsilon, y + \varepsilon) \int_{-\varepsilon}^{\varepsilon} g(1-t, y + \varepsilon, u + \varepsilon) du dy}{\int_{-\varepsilon}^{\varepsilon} g(1-s, x + \varepsilon, u + \varepsilon) du}. \end{aligned}$$

Divide numerator and denominator by ε^2 and use the same application of l'Hospital's rule as in above we get

$$\lim_{\varepsilon \rightarrow 0} \bar{p}_\varepsilon(s, x, t, y) = p_0^+(s, x, t, y),$$

which implies the convergence of the finite dimensional distributions.

The tightness of \bar{W}_ε can be shown by a slight modification of the proof of Theorem (4.1) in [2].

3. Conditioned limit theorem for weakly dependent random variables.

Let $\{X_i, i \in Z\}$ be a sequence of strictly stationary random variables. Let \mathcal{F}_n^m denote the σ -field generated by random variables $X_i, i=m, m+1, \dots, n$. Suppose that the sequence $\{X_i, i \in Z\}$ satisfies the strong mixing condition, that is

$$\alpha(n) = \sup_{A \in \mathcal{F}_{-\infty}^0, B \in \mathcal{F}_n^\infty} |P[A \cap B] - P[A]P[B]| \rightarrow 0 \text{ as } n \rightarrow \infty.$$

If the following limit exists, $\sigma^2 \equiv \lim_{n \rightarrow \infty} n^{-1} E \left\{ \sum_{i=1}^n X_i \right\}^2 > 0$, then define a random element Y_n on $C[0, 1]$ by

$$Y_n(t) = \begin{cases} (nt)X_1/(\sigma n^{1/2}), & \text{for } t \in [0, 1/n] \\ \sum_{i=1}^k X_i/(\sigma n^{1/2}) + (nt-k)X_{k+1}/(\sigma n^{1/2}), & \text{for } t \in (k/n, (k+1)/n], \\ & k=1, \dots, n-1. \end{cases}$$

Let P_n be the distribution of $\{Y_n(t), 0 \leq t \leq 1\}$ and Q the Wiener measure on (C, \mathcal{C}) . The Prokhorov-Lévy metric $\rho(\cdot, \cdot)$ on the space of probability measures on (C, \mathcal{C}) is defined by

$$\rho(R, P) = \inf \{ \varepsilon > 0; R[B] \leq \varepsilon + P[y; d(x, y) < \varepsilon, x \in B], \\ P[B] \leq \varepsilon + R[y; d(x, y) < \varepsilon, x \in B] \text{ for all } B \in \mathcal{C} \},$$

where R and P are probability measures on (C, \mathcal{C}) .

S. Kanagawa ([3] page 104) proved the following:

Theorem (1 [3] page 104). *Let $\{X_i, i \in Z\}$ be a sequence of strictly stationary random variables with $EX_1=0$ and $E|X_1|^r < \infty$ for some $r > 2$. Suppose that the sequence $\{X_i, i \in Z\}$ satisfies the strong mixing condition with coefficient $\alpha(n)$ and that there exists s with $2 < s < r$ such that*

$$(3.1) \quad \sum_{i=1}^{\infty} (\alpha(i))^{(r-s)/rs} < \infty.$$

If $s \leq 4$, then for any $\delta < s(s-2)/\{4(s-1)(s+1)\}$, we have

$$\rho(P_n, Q) = o(n^{-\delta}), \text{ as } n \rightarrow \infty.$$

If $r > 4$ and (3.1) holds for some s with $4 < s < r$, then for any $k < (s-4)/30(s+1)$, we have

$$\rho(P_n, Q) = o(n^{-2/15-k}), \text{ as } n \rightarrow \infty.$$

Define $m_n = \inf_{0 \leq t \leq 1} Y_n(t)$ and put $\bar{Y}_n = (Y_n | m_n > \varepsilon_n, Y_n(1) < \varepsilon_n)$, $n \geq 1$, where $\{\varepsilon_n, n \geq 1\}$ is a sequence of positive constants.

Now we prove the following theorem.

Theorem (3.1). Under the assumption of Theorem 1 ([3] page 104) $\bar{Y}_n \Rightarrow W_0^+$, $n \rightarrow \infty$ provided $\{\varepsilon_n, n \geq 1\}$ is a sequence of positive numbers such that, in the case $s \leq 4$, for every $\delta < s(s-2)/\{4(s-1)(s+1)\}$, $(n^\delta \varepsilon_n^2) = O(1)$ as $n \rightarrow \infty$,

and, in the case $4 < s < r$, for every $k < (s-4)/30(s+1)$, $(n^{2/15+k} \varepsilon_n^2) = O(1)$ as $n \rightarrow \infty$.

Proof. Let A be a set such that $P[W_0^+ \in \partial A] = 0$. We are going to show that

$$(3.2) \quad \lim_{n \rightarrow \infty} P[\bar{Y}_n \in A] = P[W_0^+ \in A].$$

Indeed, we have

$$\begin{aligned} P[\bar{Y}_n \in A] &= \{P[Y_n \in A, m_n > -\varepsilon_n, Y(1) < \varepsilon_n] - P[W \in A, m(1) > -\varepsilon_n, W(1) < \varepsilon_n] \\ &\quad + P[W \in A, m(1) > -\varepsilon_n, W(1) < \varepsilon_n]\} / \{P[m_n > -\varepsilon_n, Y_n(1) < \varepsilon_n] \\ &\quad - P[m(1) > -\varepsilon_n, W(1) < \varepsilon_n] + P[m(1) > -\varepsilon_n, W(1) < \varepsilon_n]\} \\ &= \{I_1^{(n)} + P[W \in A, m(1) > -\varepsilon_n, W(1) > \varepsilon_n]\} / \{I_2^{(n)} + P[m(1) > -\varepsilon_n, \\ &\quad W(1) < \varepsilon_n]\}. \end{aligned}$$

By our assumptions and Theorem 1 ([3] page 104)

$$I_1^{(n)}/\varepsilon_n^2 \rightarrow 0, \quad I_2^{(n)}/\varepsilon_n^2 \rightarrow 0, \quad \text{as } n \rightarrow \infty.$$

Furthermore, one can verify that

$$P[m(1) > -\varepsilon_n, W(1) < \varepsilon_n] / \varepsilon_n^2 \rightarrow \sigma / (2\pi)^{1/2},$$

and by Theorem (2.1) we have

$$\lim_{n \rightarrow \infty} P[W \in A \mid m(1) > -\varepsilon_n, W(1) < \varepsilon_n] = P[W_0^+ \in A],$$

so that we get (3.2). This fact proves Theorem (3.1).

References

- [1] Billingsley, P.: "Convergence of Probability Measures" Wiley, New York (1968).
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- [3] Kanagawa, S.: *Rates of convergence in the invariance principle for weakly dependent random variables*. Yokohama Math. J., 30 (1982), 103-119.

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