

# RATES OF CONVERGENCE IN THE ERDÖS-KAC TYPE INVARIANCE PRINCIPLE FOR WEAKLY DEPENDENT RANDOM VARIABLES

By

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## 1. Introduction and results.

Let  $\{X_i, i \in \mathbb{Z}\}$  be a strictly stationary sequence of random variables. Let  $\mathcal{F}_p^q$  denote the  $\sigma$ -field generated by random variables  $\{X_i, i = p, p+1, \dots, q\}$ . Suppose that the sequence  $\{X_i\}$  satisfies the condition of the absolute regularity, that is

$$(1.1) \quad \beta(n) = E \left\{ \sup_{A \in \mathcal{F}_n^\infty} |P(A | \mathcal{F}_{-\infty}^0) - P(A)| \right\} \downarrow 0,$$

as  $n \rightarrow \infty$ .

Suppose  $EX_1 = 0$  and  $E|X_1|^{2+\delta} < \infty$  for some  $\delta > 0$ . Besides if we assume

$$\sum_{i=1}^{\infty} (\beta(n))^{\delta/(2+\delta)} < \infty,$$

then the series

$$(1.2) \quad \sigma^2 \equiv EX_1^2 + 2 \sum_{i=1}^{\infty} E(X_1 X_{i+1})$$

converges absolutely. (See e.g. [8].) We suppose moreover that  $\sigma > 0$ .

Define a continuous polygonal line  $\{X_n(t), 0 \leq t \leq 1\}$  by

$$X_n(t) = (\sigma n^{1/2})^{-1} \sum_{i=1}^{[nt]} X_i + (nt - [nt])(\sigma n^{1/2})^{-1} X_{[nt]+1},$$

where  $[x]$  denotes the integral part of  $x$ . Let  $\{B(t), t \geq 0\}$  be a standard Brownian motion. Moreover let  $P_n$  be distribution of  $\{X_n(t), 0 \leq t \leq 1\}$  and  $W$  the Wiener measure on the space of continuous functions  $C[0, 1]$ . As is well known, if the above conditions are satisfied, then the invariance principle of Donsker type holds, that is, as  $n \rightarrow \infty$

$$(1.3) \quad P_n \Rightarrow W. \quad (\text{Weak convergence in } C[0, 1].)$$

(See e.g. Oodaira and Yoshihara [4], Kato [3].) Furthermore, from Theorem 5.1 in Billingsley [1], (1.3) implies the Erdős-Kac type invariance principle, that is

$$(1.4) \quad \lim_{n \rightarrow \infty} P(\max_{1 \leq k \leq n} |S_k| < \sigma n^{1/2} \lambda) = T(\lambda) = P(\sup_{0 \leq t \leq 1} |B(t)| < \lambda)$$

for all  $\lambda$ , where

$$S_k = \sum_{i=1}^k X_i$$

and

$$(1.4') \quad T(\lambda) = \frac{4}{\pi} \sum_{k=0}^{\infty} \frac{(-1)^k}{2k+1} \exp\left(-\frac{(2k+1)^2 \pi^2}{8\lambda^2}\right).$$

For the convergence rate in (1.4) Yoshihara [8] gave the following result; if  $E|X_1|^{4+\varepsilon} < \infty$  for some  $\varepsilon > 0$  and

$$(1.5) \quad \sum_{k=1}^{\infty} k(\beta(k))^{\varepsilon/(4+\varepsilon)} < \infty,$$

then

$$(1.6) \quad \Delta_n \equiv \sup_{\lambda} |P(\max_{1 \leq k \leq n} |S_k| < \sigma n^{1/2} \lambda) - T(\lambda)| = O(n^{-1/8} (\log n)^{1/2}).$$

The method of the proof in Yoshihara [8] highly depends on the assumption that  $E|X_1|^{4+\varepsilon} < \infty$  for some  $\varepsilon$ . No stronger condition on moments seems to imply better rates in (1.6). In this paper we shall give the better rate in (1.4) under stronger moment condition and furthermore we give the more results under several other moment conditions.

**Theorem 1.** Let  $\{X_i, i \in \mathbf{Z}\}$  be a sequence of strictly stationary random variables satisfying  $EX_1 = 0$  and

$$(1.7) \quad P(|X_1| \geq \lambda) \leq K_1 \exp(-\alpha \lambda^\mu)$$

for some  $\alpha > 0$  and  $\mu > 0$ , where  $K_1$  is a positive constant. Suppose that  $\{X_i\}$  satisfies the condition of the absolute regularity with coefficient  $\beta(\cdot)$  such that

$$(1.8) \quad \beta(n) = o(e^{-\gamma n})$$

for some  $\gamma > 0$ . Then we have

$$(1.9) \quad \Delta_n = O(n^{-1/4} (\log n)^\zeta),$$

where  $\zeta = \max(1+\varepsilon, 1/2+1/\mu)$  for any  $\varepsilon > 0$ .

**Theorem 2.** In Theorem 1, replace condition (1.7) by

$$(1.10) \quad E|X_1|^r < \infty \quad \text{for some } r > 6.$$

Then we have for any  $\delta < (r-2)/4(r+1)$

$$(1.11) \quad \Delta_n = o(n^{-\delta})$$

as  $n \rightarrow \infty$ .

**Theorem 3.** In Theorem 1, replace condition (1.7) by

$$(1.12) \quad E|X_1|^r < \infty \quad \text{for some } r \text{ with } 4 < r \leq 6.$$

Then we have

$$(1.13) \quad \Delta_n = o(n^{-1/7})$$

as  $n \rightarrow \infty$ . Moreover if

$$(1.14) \quad E|X_1|^r < \infty \quad \text{for some } r \text{ with } 2 < r \leq 4,$$

then we have for any  $\delta < (r-2)/2(2r-1)$ ,

$$(1.15) \quad \Delta_n = o(n^{-\delta})$$

as  $n \rightarrow \infty$ .

## 2. Preliminaries.

The basic idea of the proofs of Theorems 1-3 is to use the following lemma which is an approximation theorem given by Yoshihara [7].

**Lemma 1.** Let  $\{X_i\}$  be an absolutely regular sequence of random variables such that  $EX_1=0$  and  $E|X_1|^{2+\varepsilon} < \infty$  for some  $\varepsilon > 0$ . Let  $g(t_1, \dots, t_k)$  be any Borel function on  $\mathbf{R}^k$  with  $|g(t_1, \dots, t_k)| \leq K_2$ ,  $K_2$  being a positive constant. Then there exists a sequence of independent random variables  $\{Y_i\}$  such that each  $Y_i$  has the same distribution as that of  $X_i$  and

$$|E\{g(X_{i_1}, X_{i_2}, \dots, X_{i_k})\} - E\{g(Y_{i_1}, Y_{i_2}, \dots, Y_{i_k})\}| \leq 2K_2 k \beta(d)$$

for any  $1 \leq i_1 < i_2 < \dots < i_k$ , where

$$d = \max_{1 \leq j \leq k-1} (i_{j+1} - i_j).$$

The following lemma is also due to Yoshihara [7].

**Lemma 2.** Let  $\{X_i\}$  be a strictly stationary and absolutely regular sequence of random variables with zero mean and  $E|X_1|^s < \infty$  for some  $s \geq 3$ . If the assumption (1.6) holds and  $\sigma > 0$  in (1.2), then

$$\sup_{t > ((s-2)\log n)^{1/2}} \{t^s (\log n)^{-1} P(|S_n| \geq t(E X_1^2)^{1/2} (\log n)^{1/2} n^{1/2})\} \leq K_s n^{-(s-2)/2},$$

where  $K_s$  is a positive constant depending only on  $s$ .

We now define several sequences of random variables associated with  $\{X_i\}$  in the following way. Let  $M = [n^{1/2}] + 1$ . For  $j = 1, \dots, M-1$ , define

$$I_j = \{(j-1)[n^{1/2}] + 1, (j-1)[n^{1/2}] + 2, \dots, j[n^{1/2}]\}$$

and

$$I_M = \{(M-1)[n^{1/2}] + 1, (M-1)[n^{1/2}] + 2, \dots, n\}.$$

Let

$$y_j = \sum_{i \in I_j} n^{-1/2} X_i \quad \text{for } j=1, \dots, M.$$

Moreover define  $U_j = \{j[n^{1/2}] - [\theta \log n] + 1, j[n^{1/2}] - [\theta \log n] + 2, \dots, j[n^{1/2}]\}$  and

$$v_j = \sum_{i \in U_j} n^{-1/2} X_i \quad \text{for } j=1, \dots, M-1 \text{ and } v_M=0,$$

where  $\theta$  is some positive constant. Denote

$$\xi_j = y_j - v_j \quad \text{for } j=1, \dots, M.$$

Using Theorem 1 in Yokoyama [6] we easily have the following lemma.

**Lemma 3.** As  $n \rightarrow \infty$ ,

$$(i) \quad E y_j^2 = n^{-1/2} + O(n^{-1}), \quad j=1, \dots, M-1,$$

$$(ii) \quad E y_M^2 = (n - (M-1)[n^{1/2}])/n + O(n^{-1})$$

and

$$(iii) \quad E v_j^2 = (\theta \log n)/n + O(n^{-1}), \quad j=1, \dots, M-1.$$

For each positive integer  $m$ ,

$$(iv) \quad E |y_j|^{2m} \leq K_\beta K_m E |X_1|^{2m} n^{-m/2}, \quad j=1, \dots, M$$

and

$$(v) \quad E |v_j|^{2m} \leq K_\beta K_m E |X_1|^{2m} (\theta \log n)^m n^{-m}, \quad j=1, \dots, M-1,$$

where  $K_m = 2m!(2m-1)$  and  $K_\beta$  is a positive constant depending only on the coefficient  $\beta(\cdot)$ .

Define  $M+1$  points  $\{a_k\}$  on  $[0, 1]$  by  $a_k = k[n^{1/2}]/n$  for  $k=0, 1, \dots, M-1$  and  $a_M=1$ . Let  $\{X_n(t), 0 \leq t \leq 1\}$ ,  $\{\tilde{X}_n(t), 0 \leq t \leq 1\}$  and  $\{\hat{X}_n(t), 0 \leq t \leq 1\}$  be continuous polygonal lines defined respectively by

$$X_n(t) = \begin{cases} n^{1/2} t X_1, & \text{for } t \in [0, 1/n], \\ n^{-1/2} S_k + (nt - k) n^{-1/2} X_{k+1}, & \text{for } t \in (k/n, (k+1)/n], k=1, \dots, M-1, \end{cases}$$

$$\tilde{X}_n(t) = \begin{cases} \frac{t}{a_1} y_1, & \text{for } t \in [0, a_1], \\ \sum_{i=1}^k y_i + \frac{t - a_k}{a_{k+1} - a_k} y_{k+1}, & \text{for } t \in (a_k, a_{k+1}], k=1, \dots, M-1, \end{cases}$$

and

$$\hat{X}_n(t) = \begin{cases} \frac{t}{a_1} y_1, & \text{for } t \in [0, a_1], \\ \sum_{i=1}^k \xi_i + \frac{t - a_k}{a_{k+1} - a_k} \xi_{k+1}, & \text{for } t \in (a_k, a_{k+1}], k=1, \dots, M-1. \end{cases}$$

### 3. More lemmas.

We need more lemmas. Let  $\varepsilon_n = n^{-1/4}(\log n)^\zeta$ . Define

$$\xi_j^* = \begin{cases} \xi_j & \text{for } |\xi_j| \leq \varepsilon_n \\ 0 & \text{for } |\xi_j| > \varepsilon_n, \end{cases}$$

and

$$\eta_j = \xi_j^* - E\xi_j^* \quad \text{for } j=1, \dots, M.$$

Moreover define continuous polygonal lines  $\{X_n^*(t), 0 \leq t \leq 1\}$  and  $\{X_n^{**}(t), 0 \leq t \leq 1\}$  by

$$X_n^*(t) = \begin{cases} \frac{t}{a_1} \xi_1^*, & \text{for } t \in [0, a_1], \\ \sum_{i=1}^k \xi_i^* + \frac{t-a_k}{a_{k+1}-a_k} \xi_{k+1}^*, & \text{for } t \in (a_k, a_{k+1}], k=1, \dots, M-1, \end{cases}$$

and

$$X_n^{**}(t) = \begin{cases} \frac{t}{a_1} \eta_1, & \text{for } t \in [0, a_1], \\ \sum_{i=1}^k \eta_i + \frac{t-a_k}{a_{k+1}-a_k} \eta_{k+1}, & \text{for } t \in (a_k, a_{k+1}], k=1, \dots, M-1. \end{cases}$$

The expectations of  $\xi_j^*$  are estimated as follows.

**Lemma 4.** As  $n \rightarrow \infty$ , for any  $u > 1/4$ ,

$$|E\xi_j^*| \leq K_u n^{-u} \quad \text{for } j=1, \dots, M,$$

where  $K_u$  is a positive constant depending only on  $u$ .

**Proof.** Since  $E\xi_j = 0$  for each  $j$ ,

$$(3.1) \quad |E\xi_1^*| = \left| \int_{|x| > \varepsilon_n} x P\left(n^{-1/2} \sum_{i=1}^{[n^{1/2}] - [\theta \log n]} X_i \in dx\right) \right| \\ \leq \left| \int_{|x| > (\log n)^\zeta} x P\left(n^{-1/4} \sum_{i=1}^{[n^{1/2}] - [\theta \log n]} X_i \in dx\right) \right|.$$

On the other hand we have by Lemma 2 that

$$(3.2) \quad P\left(\left|n^{-1/4} \sum_{i=1}^{[n^{1/2}] - [\theta \log n]} X_i\right| \geq x\right) \leq K_s x^{-s} n^{-(s-2)/4} (\log n)^{1+s/2},$$

for  $x > (\log n)^\zeta$ . Thus, from (3.1) and (3.2), the lemma is proved.

**Lemma 5.** As  $n \rightarrow \infty$ ,

$$P\left(\sup_{0 \leq t \leq 1} |X_n(t) - \hat{X}_n(t)| \geq \varepsilon_n\right) = o(\varepsilon_n)$$

**Proof.** We see from the definitions of  $X_n(t)$  and  $\hat{X}_n(t)$  that

$$\begin{aligned}
(3.3) \quad & P\left(\sup_{0 \leq t \leq 1} |X_n(t) - \hat{X}_n(t)| \geq \varepsilon_n\right) \\
& \leq P\left(\sup_{0 \leq t \leq 1} |X_n(t) - \tilde{X}_n(t)| \geq \varepsilon_n/2\right) + P\left(\sup_{0 \leq t \leq 1} |\tilde{X}_n(t) - \hat{X}_n(t)| \geq \varepsilon_n/2\right) \\
& \leq \sum_{i=1}^M P\left(\sup_{\alpha_{i-1} \leq t \leq \alpha_i} |X_n(t) - \hat{X}_n(t)| \geq \varepsilon_n/2\right) + P\left(\max_{1 \leq k \leq M} \left|\sum_{i=1}^k v_i\right| \geq \varepsilon_n/2\right) \\
& \leq MP\left(\max_{1 \leq k \leq [n^{1/2}]} \left|\sum_{i=1}^k n^{-1/2} X_i\right| \geq \varepsilon_n/2\right) + \sum_{k=1}^M P\left(\left|\sum_{i=1}^k v_i\right| \geq \varepsilon_n/2\right) \\
& \equiv A_1 + A_2, \quad \text{say.}
\end{aligned}$$

Recall that  $\zeta > 1$ . Using Lemma 2 with  $s=8$ , we have that

$$\begin{aligned}
(3.4) \quad & P\left(\left|\sum_{i=1}^k n^{-1/2} X_i\right| \geq \varepsilon_n/2\right) = P\left(\left|\sum_{i=1}^k n^{-1/4} X_i\right| \geq (\log n)^{\zeta/2}\right) \\
& = o(n^{-3/2}),
\end{aligned}$$

uniformly in  $k$  with  $1 \leq k \leq [n^{1/2}]$ . Hence

$$(3.5) \quad A_1 \leq M \sum_{k=1}^{[n^{1/2}]} P\left(\left|\sum_{i=1}^k n^{-1/2} X_i\right| \geq \varepsilon_n/2\right) = o(\varepsilon_n).$$

As to  $A_2$  we also apply Lemma 2 to see that

$$\begin{aligned}
(3.6) \quad & A_2 \leq \sum_{k=1}^M P\left(\left|\sum_{i=1}^k v_i\right| \geq \varepsilon_n/2\right) \\
& \leq \sum_{k=1}^M P\left(\left|\sum_{i=1}^k v_i\right| \geq \theta^{-1/2} (E v_i^2)^{1/2} (\log n)^{\zeta-1/2} n^{1/4}\right) \\
& = o(\varepsilon_n).
\end{aligned}$$

Thus the lemma is proved from the relations (3.3), (3.5) and (3.6).

**Lemma 6.** As  $n \rightarrow \infty$ ,

$$P\left(\sup_{0 \leq t \leq 1} |\hat{X}_n(t) - X_n^{**}(t)| \geq \varepsilon_n\right) = o(\varepsilon_n).$$

**Proof.** From the definitions of  $X_n(t)$  and  $X_n^{**}(t)$ , we see that

$$\begin{aligned}
(3.7) \quad & P\left(\sup_{0 \leq t \leq 1} |\hat{X}_n(t) - X_n^{**}(t)| \geq \varepsilon_n\right) \\
& \leq P\left(\sup_{0 \leq t \leq 1} |\hat{X}_n(t) - X_n^*(t)| \geq \varepsilon_n/2\right) + P\left(\sup_{0 \leq t \leq 1} |X_n^*(t) - X_n^{**}(t)| \geq \varepsilon_n/2\right) \\
& = P\left(\max_{1 \leq k \leq M} \left|\sum_{i=1}^k (\xi_i - \xi_i^*)\right| \geq \varepsilon_n/2\right) + P\left(\max_{1 \leq k \leq M} \left|\sum_{i=1}^k E \xi_i^*\right| \geq \varepsilon_n/2\right) \\
& \equiv C_1 + C_2, \quad \text{say.}
\end{aligned}$$

Since  $\xi_i = \xi_i^*$  if  $|\xi_i| \leq \varepsilon_n$  for each  $i$ , we have from (3.4) that

$$(3.8) \quad C_1 \leq P(\max_{1 \leq k \leq M} |\xi_k| \geq \varepsilon_n) \leq MP(|\xi_1| \geq \varepsilon_n) = o(\varepsilon_n).$$

Moreover, as to  $C_2$ , since we have from Lemma 3 that

$$\max_{1 \leq k \leq M} \left| \sum_{i=1}^k E \xi_i^* \right| \leq \sum_{i=1}^M |E \xi_i^*| = o(\varepsilon_n),$$

it follows that for sufficiently large  $n$

$$(3.9) \quad C_2 \equiv 0.$$

Therefore we obtain the lemma from the relations (3.7)–(3.9).

#### 4. Proof of Theorem 1.

We now turn out to prove Theorem 1. Define the distribution function  $F_n^*(\cdot)$  as follows,

$$F_n^*(\lambda) = P\left(\max_{1 \leq k \leq M} \left| \sum_{i=1}^k \eta_i \right| \leq \lambda\right).$$

Furthermore denote  $\mathcal{M}_p^q$  the  $\sigma$ -field generated by the random variables  $\{\eta_p, \eta_{p+1}, \dots, \eta_q\}$  for all  $p$  and  $q$  with  $1 \leq p \leq q \leq M$ . We have from the assumption (1.1) that

$$\max_{1 \leq k \leq M-1} E \left\{ \sup_{A \in \mathcal{M}_{k+1}^M} |P(A | \mathcal{M}_1^k) - P(A)| \right\} = O(n^{-\tau\theta}).$$

Therefore, using Lemma 1, we see that there exists a sequence of independent random variables  $\{Y_1, \dots, Y_M\}$  such that  $Y_i$  has the same distribution as that of  $\eta_i$  for each  $i$  and

$$(4.1) \quad \left| P\left(\max_{1 \leq k \leq M} \left| \sum_{i=1}^k \eta_i \right| \geq \lambda\right) - P\left(\max_{1 \leq k \leq M} \left| \sum_{i=1}^k Y_i \right| \geq \lambda\right) \right| = O(n^{-\tau\theta+1/2})$$

for any  $\lambda$ .

On the other hand, applying the Skorokhod embedding theorem (see e.g. [5]) to  $\{Y_i\}$ , we can construct a Brownian motion  $\{B(t), t \geq 0\}$  and a sequence of independent and positive random variables  $\{T_i, 1 \leq i \leq M\}$  such that the joint distributions of

$$\left\{ B(T_1), B(T_1+T_2)-B(T_1), \dots, B\left(\sum_{i=1}^M T_i\right) - B\left(\sum_{i=1}^{M-1} T_i\right) \right\}$$

are same as those of  $\{Y_1, Y_2, \dots, Y_M\}$  and

$$(4.2) \quad ET_i = EY_i^2 \quad \text{and} \quad E|T_i|^m \leq 2m! E|Y_i|^{2m}, \quad i=1, \dots, M,$$

for each positive integer  $m$ .

Combining this argument and the equation (4.1), we have that

$$(4.3) \quad F_n^*(\lambda) = P\left(\max_{1 \leq k \leq M} \left| \sum_{i=1}^k Y_i \right| \leq \lambda\right) + o(\varepsilon_n)$$

$$=P\left(\max_{1 \leq k \leq M} \left| B\left(\sum_{i=1}^k T_i\right) \right| \leq \lambda\right) + o(\varepsilon_n).$$

Now since  $E\eta_i=0$  and  $|\eta_i| \leq 2\varepsilon_n$  for each  $i$ ,

$$\max_{1 \leq k \leq M} \sup_{0 \leq t \leq T_k} \left| B\left(t + \sum_{i=1}^{k-1} T_i\right) - B\left(\sum_{i=1}^{k-1} T_i\right) \right| \leq 2\varepsilon_n.$$

Therefore we have from (4.3) that

$$(4.4) \quad F_n^*(\lambda) \leq P\left(\sup_{0 \leq t \leq \sum_{i=1}^M T_i} |B(t)| \leq \lambda + 2\varepsilon_n\right) + o(\varepsilon_n) \\ \leq P\left(\sup_{0 \leq t \leq 1 - \varepsilon_n} |B(t)| \leq \lambda + 2\varepsilon_n\right) + P\left(\left|\sum_{i=1}^M T_i - 1\right| \geq \varepsilon_n\right) + o(\varepsilon_n),$$

and

$$(4.5) \quad F_n^*(\lambda) \geq P\left(\sup_{0 \leq t \leq 1 + \varepsilon_n} |B(t)| \leq \lambda - 2\varepsilon_n\right) - P\left(\left|\sum_{i=1}^M T_i - 1\right| \geq \varepsilon_n\right) - o(\varepsilon_n).$$

As to the right hand sides of (4.4) and (4.5), the same method as in the proof of Theorem 1 stated in Sawyer [5] gives us in view of the inequality (4.2) that

$$(4.6) \quad P\left(\left|\sum_{i=1}^M T_i - 1\right| \geq \varepsilon_n\right) = o(\varepsilon_n).$$

On the other hand we see from Lemmas 5 and 6 that

$$(4.7) \quad F_n(\lambda) = P\left(\sup_{0 \leq t \leq 1} |X_n(t)| \leq \lambda\right) \\ \leq P\left(\sup_{0 \leq t \leq 1} |X_n(t)| \leq \lambda, \sup_{0 \leq t \leq 1} |X_n(t) - X_n^{**}(t)| \leq 2\varepsilon_n\right) \\ + P\left(\sup_{0 \leq t \leq 1} |X_n(t) - \hat{X}_n(t)| \geq \varepsilon_n\right) + P\left(\sup_{0 \leq t \leq 1} |\hat{X}_n(t) - X_n^{**}(t)| \geq \varepsilon_n\right) \\ \leq F^*(\lambda + 2\varepsilon_n) + o(\varepsilon_n)$$

and

$$(4.8) \quad F_n(\lambda) \geq F^*(\lambda - 2\varepsilon_n) - o(\varepsilon_n).$$

Thus we have from the inequalities (4.4)-(4.8) that

$$P\left(\sup_{0 \leq t \leq 1 + \varepsilon_n} |B(t)| \leq \lambda - 4\varepsilon_n\right) - o(\varepsilon_n) \\ \leq F_n(\lambda) \leq P\left(\sup_{0 \leq t \leq 1 - \varepsilon_n} |B(t)| \leq \lambda + 4\varepsilon_n\right) + o(\varepsilon_n).$$

The rest of the proof is the same as that of Theorem 1 in [5] and is omitted.

## 5. Proof of Theorem 2.

By the truncation argument which was used in the proof of Theorem 2 in [5], we can easily prove the theorem.



## 6. Proof of Theorem 3.

In the proof of Theorem 1 in [2], replace Lemma 3 in [2] by Lemma 1, then we can obtain Theorem 3.

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## References

- [1] Billingsley, P.: *Convergence of Probability Measures*, Wiley, New York (1968).
- [2] Kanagawa, S.: *On the rate of convergence of the invariance principle for stationary sequences*, Keio Sci. Tech. Rep., 35 (1982), 53-61.
- [3] Kato, Y.: *Functional central limit theorem for stationary processes*, Keio Engin. Rep., 28 (1975), 23-39.
- [4] Oodaira, H. and Yoshihara, K.: *Functional central limit theorem for strictly stationary processes satisfying the strong mixing condition*, Kodai Math. Sem. Rep., 24 (1972), 259-269.
- [5] Sawyer, S.: *A uniform rate of convergence for the maximum absolute value of partial sums in probability*, Comm. Pure Appl. Math., 20 (1967), 647-658.
- [6] Yokoyama, R.: *Moments bounds for stationary mixing sequences*, Z. Wahrscheinlichkeitstheorie verw. Gebiete., 52 (1980), 45-57.
- [7] Yoshihara, K.: *Probability inequalities for sums of absolutely regular processes and their applications*, Z. Wahrscheinlichkeitstheorie verw. Gebiete., 43 (1978), 319-330.
- [8] Yoshihara, K.: *Convergence rates of the invariance principle for absolutely regular sequences*, Yokohama Math. J., 27 (1979), 49-55.

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