

## A SUPPLEMENT TO CONVERGENCE RATES IN THE LAW OF LARGE NUMBERS

By

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### 1. Introduction

Let  $\{X_n\}_{n \geq 1}$  be a sequence of i.i.d. random variables with common distribution  $F$  and let  $S_n = X_1 + \cdots + X_n$ ,  $S_0 = 0$ . Throughout this paper we suppose that the following weak law of large numbers holds for some  $\alpha$ ,  $0 < \alpha < 2$ ,

$$(1) \quad \lim_n P(|S_n| > n^{1/\alpha} \varepsilon) = 0 \quad \text{for every } \varepsilon > 0.$$

As was shown in [3] and [5], (1) holds if and only if the following two conditions are satisfied:

$$(2) \quad P(|X| > n^{1/\alpha}) = o(n^{-1}),$$

where  $X = X_1$ , and

$$(3) \quad \int_{|x| > n^{1/\alpha}} xF(dx) = o(n^{1/\alpha - 1}).$$

Note that if  $0 < \alpha < 1$  then (2) implies (3), if  $1 \leq \alpha < 2$  then (2) and  $E(X) = 0$  are sufficient for (3).

It is well-known ([1], [5], [6]) that

$$(4) \quad \sum_{n=1}^{\infty} P(|S_n| > n^{1/\alpha} \varepsilon) < \infty \quad \text{for every } \varepsilon > 0,$$

if and only if  $E|X|^{2\alpha} < \infty$  and (3) hold. If (4) holds then

$$(5) \quad \lim_{x \rightarrow \infty} \sum_{0 \leq n \leq (x/\varepsilon)^\alpha} P(x \leq S_n \leq x + h) = 0$$

for every  $\varepsilon > 0$  and  $h > 0$ . However as will be shown later (1) does not imply (5). In this note we give a sufficient condition for (5), and show that it is close to the best possible one. Note that (5) can be written as follows:

$$(6) \quad \lim_{x \rightarrow \infty} EN((x/\varepsilon)^\alpha, [x, x + h]) = 0,$$

where  $N(x, I)$  is the number of  $S_k$ ,  $0 \leq k \leq x$ , falling within an interval  $I$ .

This paper stemmed from an attempt to improve Bickel and Yahav's renewal

theorem in the plane [2]. Let  $\|\cdot\|$  be a norm in  $R^2$  such that its unit ball  $B = \{x; \|x\| \leq 1\}$  is a convex polygon. Let  $\{X_n\}$  be a sequence of i.i.d. two-dimensional random vectors having finite non-zero expectation, and let  $S_n = X_1 + \cdots + X_n$ . Suppose that the distribution of  $X_1$  is not supported by a proper closed subgroup of  $R^2$ . Bickel and Yahav showed that if either  $E|X_1|^2 < \infty$  or  $E(X_1)$  is not parallel to sides of  $B$  then the following Blackwell type renewal theorem:

$$\lim_{x \rightarrow \infty} \sum_{n=1}^{\infty} P(x \leq \|S_n\| \leq x+h) = h/\|E(X_1)\|$$

holds for  $h > 0$ . Our theorem enables us to show that when  $E(X_1)$  is parallel to a side of  $B$  the condition  $E|X_1|^2 < \infty$  cannot be weakened to  $E|X_1|^r < \infty, r < 3/2$ .

## 2. Results

The main result we are going to prove is the following:

**Theorem.** (i) *If (3) and*

$$(7) \quad P(|X| > n) = o(n^{-3\alpha/2})$$

*are satisfied then (5) holds for every  $\varepsilon > 0$  and  $h > 0$ .*

(ii) *In the preceding statement condition (7) cannot be replaced by*

$$(8) \quad P(|X| > n) = o(n^{-3\alpha/2}\psi(n)),$$

*where  $\psi(x)$  is an arbitrary nonnegative function on  $[0, \infty)$  which increases to  $+\infty$ .*

Note that a random walk which satisfies (1) and fails to satisfy (5) must be recurrent. In fact suppose  $\{S_n\}$  is transient. If  $E|X| = \infty$  then

$$\lim_{x \rightarrow \infty} \sum_{n=1}^{\infty} P(x \leq S_n \leq x+h) = 0$$

(see [4], p. 368) and therefore  $\{S_n\}$  satisfies (5). If  $E|X| < \infty$  then we must have  $E(X) \neq 0$  since  $\{S_n\}$  is transient. In this case (1) does not hold  $1 \leq \alpha < 2$ , and both (1) and (5) hold if  $0 < \alpha < 1$ .

As is well-known  $E|X|^{3\alpha/2} < \infty$  implies (7). Furthermore if

$$P(|X| > n) \sim n^{-3\alpha/2} \log n,$$

then  $E|X|^{3\alpha/2} = \infty$ , while  $E|X|^r < \infty$  for  $r < 3\alpha/2$ . Therefore we obtain the following:

**Corollary.** (i) *If  $0 < \alpha < 1$  then  $E|X|^{3\alpha/2} < \infty$  implies (5). If  $1 \leq \alpha < 2$  then  $E|X|^{3\alpha/2} < \infty$  and  $E(X) = 0$  imply (5).*

(ii) *In the above statements the condition  $E|X|^{3\alpha/2} < \infty$  cannot be replaced by the following:  $E|X|^r < \infty$  for every  $r < 3\alpha/2$ .*

*Proof of Theorem.*

(i) It is obvious that if (5) holds for  $h=1$  then it holds every finite  $h > 0$ . Our

method of proof does not depend on the choice of  $\varepsilon > 0$ . Therefore we may assume  $h=1$  and  $\varepsilon=1$ . Furthermore it suffices to prove (5) for nondegenerate  $X$ , since the statement (i) is obvious for degenerate random variables. For nondegenerate case we have (see [7], p. 49)

$$(9) \quad \sup_x P(x \leq S_n \leq x+1) \leq Cn^{-1/2},$$

where  $C$  is a constant depending only on the distribution of  $X$ . Therefore we obtain

$$(10) \quad \sup_y EN(x^\alpha, [y, y+1]) \leq C \sum_{n \leq x^\alpha} n^{-1/2} = O(x^{\alpha/2})$$

as  $x \rightarrow \infty$ . The following inequality is also well-known ([7], p. 50)

$$(11) \quad P\left(\max_{1 \leq k \leq n} S_k - m(S_k - S_n) \geq x\right) \leq 2P(S_n \geq x),$$

where  $m(X)$  is a median of a random variable  $X$ . From (1) we have  $m(S_n) = o(n^{1/\alpha})$ , and therefore for large  $n$

$$(12) \quad P\left(\max_{1 \leq k \leq n} S_k \geq n^{1/\alpha}\right) \leq 2P(S_n > 2^{-1}n^{1/\alpha}).$$

It follows from a theorem of Heyde-Rohatgi [5] that (3) and (7) implies

$$(13) \quad P(|S_n| > 2^{-1}n^{1/\alpha}) = o(n^{-1/2}).$$

From (12) and (13) we obtain

$$(14) \quad P\left(\max_{1 \leq k \leq n^\alpha} S_k \geq n\right) = o(n^{-\alpha/2}).$$

Let  $T_K$  denote the first time the random walk  $\{S_n\}$  hits a set  $K$ :

$$T_K = \inf\{k \geq 0; S_k \in K\},$$

and  $T_K = \infty$  if  $S_k \notin K$  for every  $k$ . In what follows we write for simplicity  $I = [0, 1]$  and  $J = [-1, 1]$ . Then we have

$$\begin{aligned} EN(x^\alpha, I+x) &= \sum_{k=0}^{[x^\alpha]} E[N(x^\alpha, I+x), T_{I+x}=k] \\ &= \sum_{k=0}^{[x^\alpha]} P(T_{I+x}=k) \cdot E[N(x^\alpha, I+x) | T_{I+x}=k] \\ &\leq \sum_{k=0}^{[x^\alpha]} P(T_{I+x}=k) \cdot EN(x^\alpha - k, J) \\ &\leq EN(x^\alpha, J) \cdot P(S_k \in I+x \text{ for some } k \leq x^\alpha) \\ &\leq EN(x^\alpha, J) \cdot P\left(\max_{1 \leq k \leq n^\alpha} S_k \geq x\right). \end{aligned}$$

It follows from (10) and (14) that this tends to zero as  $x \rightarrow \infty$ .

(ii) We shall construct a sequence  $\{X_n\}$  of i.i.d. symmetric random variables which satisfies (8) and fails to satisfy (6). Let  $0 < \psi(x) \uparrow \infty$ . Choose  $\phi(x)$  satisfying  $0 < \phi(x) \uparrow \infty$ ,  $x^{-\alpha/2} \phi(x) \downarrow 0$ , and  $\phi(x)/\psi(x) \rightarrow 0$  as  $x \rightarrow \infty$ . Let  $\{X_n^{(j)}\}_{n \geq 1, j \geq 0}$  be a double sequence of mutually independent random variables such that  $X_1^{(j)}, X_2^{(j)}, \dots$  have a common distribution for every  $j \geq 0$ . Let  $X_n^{(0)}$  be uniformly distributed over  $[-1, 1]$ , and for  $j \geq 1$  let

$$\begin{aligned} P(X_n^{(j)} = a_j) &= P(X_n^{(j)} = -a_j) = p_j, \\ P(X_n^{(j)} = 0) &= 1 - 2p_j, \end{aligned}$$

where  $a_j > 0$  and  $p_j > 0$  are determined later.

For every  $k \geq 1$ ,

$$S_n^{(k-1)} = \sum_{m=1}^n \sum_{j=0}^{k-1} X_m^{(j)}$$

is a sum of  $n$  i.i.d. random variables, each having an absolutely continuous distribution with zero mean and finite variance. Hence by Shepp's local limit theorem (see [7], p. 214) we can find for every  $k \geq 1$  two constants  $c_k > 0$  and  $n_k > 0$  such that

$$(15) \quad P(0 \leq S_n^{(k-1)} \leq 1) \geq 2c_k n^{-1/2} \quad \text{for } n \geq n_k.$$

By (15) we can choose an increasing sequence  $\{a_n\}$ ,  $a_n > 0$ , such that

$$(16) \quad \sum_{j=1}^{\infty} a_j^{-\alpha/2} \phi(a_j) < \infty,$$

$$(17) \quad 2^{3\alpha/2} a_{j+1}^{-3\alpha/2} \phi(a_{j+1}) \leq a_j^{-3\alpha/2} \phi(a_j),$$

and

$$(18) \quad \sum_{n=l_k}^{m_k} P(0 \leq S_n^{(k-1)} \leq 1) \geq a_k^{\alpha/2} / \phi(a_k) \quad \text{for } k \geq 1,$$

where  $l_k = [a_k^\alpha/2]$  and  $m_k = [a_k^\alpha]$ . Let

$$p_j = \min(a_j^{-3\alpha/2} \phi(a_j), 1/2).$$

Then we have by (16)

$$\sum_{j=1}^{\infty} P(|X_1^{(j)}| \neq 0) = 2 \sum_{j=1}^{\infty} p_j < \infty.$$

Hence by Borel-Cantelli lemma  $\sum_{j=1}^{\infty} X_n^{(j)}$  converges a.s. to a random variable  $X_n$  for each  $n \geq 1$ . Thus we obtain a sequence  $\{X_n\}$  of i.i.d. random variables satisfying (3). From (17) we have  $2a_{j-1} < a_j$  and therefore  $a_1 + \dots + a_{j-1} < a_j$  for  $j \geq 2$ . Hence it follows from (17) that for  $2a_{j-1} \leq x \leq 2a_j$

$$\begin{aligned}
(19) \quad P(|X_1| > x) &\leq P(X_1^{(k)} \neq 0 \text{ for some } k \geq j) \\
&\leq 2 \sum_{k=j}^{\infty} p_k \leq 2 \sum_{k=j}^{\infty} a_k^{-3\alpha/2} \phi(a_k) \\
&\leq 2 \sum_{i=0}^{\infty} (2^{-3\alpha/2})^i a_j^{-3\alpha/2} \phi(a_j) \\
&\leq 2(1 - 2^{-3\alpha/2})^{-1} 2^{3\alpha/2} x^{-3\alpha/2} \phi(x).
\end{aligned}$$

The last inequality follows from the fact that  $\phi(x)$  is increasing and  $x^{-3\alpha/2}\phi(x)$  is decreasing. (19) shows that  $X_1$  satisfies (8).

It remains to prove that  $S_n = X_1 + \cdots + X_n$  does not satisfy (6). Again we suppose that  $h=1$  and  $\varepsilon=1$ . Let

$$\begin{aligned}
N_k &= N(a_k^\alpha, [a_k, a_k + 1]) \\
&= \text{no. of } n \leq m_k \text{ such that } 0 \leq S_n - a_k \leq 1, \\
N_k' &= \text{no. of } n \leq m_k \text{ such that } 0 \leq S_n^{(k)} - a_k \leq 1, \\
N_k'' &= \text{no. of } n \text{ such that } l_k \leq n \leq m_k \text{ and } 0 \leq S_n^{(k-1)} \leq 1, \\
A_k &= \{X_m^{(j)} = 0 \text{ for every } 1 \leq m \leq m_k, j \geq k+1\}, \\
B_k &= \{\text{there exists only one } m \leq l_k \text{ satisfying } X_m^{(k)} = a_k, \text{ and } X_l^{(k)} = 0 \text{ for every} \\
&\quad l \neq m, 1 \leq l \leq m_k\}.
\end{aligned}$$

Then we obtain

$$\begin{aligned}
(20) \quad EN_k &\geq EN_k \cdot 1_{A_k} = EN_k' \cdot 1_{A_k} = EN_k' \cdot P(A_k) \geq EN_k' \cdot 1_{A_k B_k} \cdot P(A_k) \\
&\geq EN_k'' \cdot 1_{A_k B_k} \cdot P(A_k) \geq EN_k'' \cdot P(A_k)^2 P(B_k).
\end{aligned}$$

It follows from (16) that

$$m_k \sum_{j=k+1}^{\infty} p_j \leq a_k^\alpha \sum_{j=k+1}^{\infty} a_j^{-3\alpha/2} \phi(a_j) \leq \sum_{j=k+1}^{\infty} a_j^{-\alpha/2} \phi(a_j) \rightarrow 0$$

as  $k \rightarrow \infty$ , and therefore

$$(21) \quad P(A_k) = \prod_{j=k+1}^{\infty} (1 - 2p_j)^{m_k} \rightarrow 1 \quad \text{as } k \rightarrow \infty.$$

It is easy to see that

$$\begin{aligned}
(22) \quad P(B_k) &\sim 2^{-1} m_k p_k (1 - 2p_k)^{m_k} \\
&\sim 2^{-1} a_k^\alpha a_k^{-3\alpha/2} \phi(a_k) [1 - 2a_k^{-3\alpha/2} \phi(a_k)]^{m_k} \\
&\sim 2^{-1} a_k^{-\alpha/2} \phi(a_k).
\end{aligned}$$

At last we have from (18) that

$$\begin{aligned}
 (23) \quad EN_k'' &= \sum_{n=l_k}^{m_k} P(0 \leq S_n^{(k-1)} \leq 1) \\
 &= \sum_{n=l_k}^{m_k} 2^{-1} P(|S_n^{(k-1)}| \leq 1) \\
 &\geq 2^{-1} a_k^{\alpha/2} / \phi(a_k),
 \end{aligned}$$

for  $k \geq 1$ . From (20), (21), (22), and (23) we have

$$\limsup_{k \rightarrow \infty} EN_k \geq 1/4.$$

Therefore

$$\limsup_{x \rightarrow \infty} EN(x^\alpha, [x, x+1]) \geq 1/4.$$

This completes the proof.

*Remark.* We can modify the above construction to show the existence of  $\{X_n\}$  satisfying (8) and

$$\limsup_{x \rightarrow \infty} EN(x^\alpha, [x, x+h]) = \infty.$$

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