

RATES OF CONVERGENCE IN THE INVARIANCE PRINCIPLE FOR WEAKLY DEPENDENT RANDOM VARIABLES

By

SHŪYA KANAGAWA

(Received April 16, 1982)

1. Introduction and results

Let $\{X_i, i \in \mathbf{Z}\}$ be a sequence of strictly stationary random variables. Let \mathcal{F}_m^n denote σ -field generated by random variables $\{X_i, i = m, m+1, \dots, n\}$. Suppose that the sequence $\{X_i\}$ satisfies the strong mixing condition, that is,

$$\alpha(n) \equiv \sup_{A \in \mathcal{F}_{-\infty}^0, B \in \mathcal{F}_n^\infty} |P(AB) - P(A)P(B)| \downarrow 0$$

as $n \rightarrow \infty$.

Suppose $EX_1 = 0$ and $E|X_i|^r < \infty$ for some $r > 2$. Under these assumptions, if

$$\sum_{i=1}^{\infty} (\alpha(i))^{(r-2)/r} < \infty,$$

then

$$(1.1) \quad \sigma^2 \equiv EX_1^2 + 2 \sum_{i=1}^{\infty} E(X_1 X_{i+1}) < \infty.$$

(See e.g. [5].) Here suppose that $\sigma > 0$.

Define a continuous polygonal line $\{X_n(t), 0 \leq t \leq 1\}$ by

$$X_n(t) = \begin{cases} (nt)X_1/(\sigma n^{1/2}), & \text{for } t \in [0, 1/n], \\ \sum_{i=1}^k X_i/(\sigma n^{1/2}) + (nt-k)X_{k+1}/(\sigma n^{1/2}), & \end{cases}$$

$$\text{for } t \in (k/n), (k+1)/n], \quad k=1, \dots, n-1.$$

Let $C = C[0, 1]$ be the space of all continuous functions on $[0, 1]$ with the uniform metric defined by

$$d(x, y) = \sup_{0 \leq t \leq 1} |x(t) - y(t)|, \quad \text{for } x, y \in C,$$

and let \mathcal{C} be the smallest σ -field containing all open sets in C . Let P_n be distribution of $\{X_n(t)\}$ and W the Wiener measure on (C, \mathcal{C}) . The Prokhorov-Lévy metric $\rho(\cdot, \cdot)$ on

the space of probability measures on (C, \mathcal{C}) is defined by

$$\begin{aligned} \rho(R, Q) &= \inf \{ \varepsilon > 0; R(B) \leq \varepsilon + Q\{y; d(x, y) < \varepsilon, x \in B\}, \\ Q(B) &\leq \varepsilon + R\{y; d(x, y) < \varepsilon, x \in B\} \text{ for all } B \in \mathcal{C} \}, \end{aligned}$$

where R and Q are probability measures on (C, \mathcal{C}) .

In this paper we shall show the following results concerning the rate of convergence of $\rho(P_n, W)$ to zero.

Theorem 1. *Let $\{X_i, i \in \mathbf{Z}\}$ be a sequence of strictly stationary random variables with $EX_1 = 0$ and $E|X_1|^r < \infty$ for some $r > 2$. Suppose that the sequence $\{X_i\}$ satisfies the strong mixing condition with coefficient $\alpha(n)$ and that there exists s with $2 < s < r$ such that*

$$(1.2) \quad \sum_{i=1}^{\infty} (\alpha(i))^{(r-s)/rs} < \infty.$$

Then (1.1) holds. Suppose that $\sigma > 0$, (and furthermore suppose $\sigma = 1$ without loss of generality).

If $s \leq 4$, then we have for any $\delta < s(s-2)/4(s-1)(s+1)$,

$$\rho(P_n, W) = o(n^{-\delta})$$

as $n \rightarrow \infty$.

If $r > 4$ and (1.2) holds for some s with $4 < s < r$, then we have for any $\kappa < (s-4)/30(s+1)$,

$$\rho(P_n, W) = o(n^{-2/15 - \kappa})$$

as $n \rightarrow \infty$.

Corollary. *In Theorem 1, suppose that*

$$\alpha(n) = O(e^{-\gamma n})$$

for some $\gamma > 0$ in place of (1.2). If $r \leq 4$, then we have for any $\delta < r(r-2)/4(r-1)(r+1)$

$$\rho(P_n, W) = o(n^{-\delta})$$

as $n \rightarrow \infty$.

If $r > 4$, then we have for any $\kappa < (r-4)/30(r+1)$

$$\rho(P_n, W) = o(n^{-2/15 - \kappa})$$

as $n \rightarrow \infty$.

It should be mentioned that Yoshihara [10] gave the rate of convergence for the stationary sequence satisfying the absolutely regular condition which is,

$$\beta(n) \equiv E \left\{ \sup_{A \in \mathcal{F}_n^\infty} |P(A | \mathcal{F}_{-\infty}^0) - P(A)| \right\} \downarrow 0$$

as $n \rightarrow \infty$.

Theorem A (Yoshihara [10]). Let $\{X_i\}$ be a strictly stationary and absolutely regular sequence of random variables with $EX_1 = 0$. If, for some $\varepsilon > 0$, $E|X_1|^{4+\varepsilon} < \infty$ and

$$(1.3) \quad \sum_{i=1}^{\infty} i(\beta(i))^{\varepsilon/(4+\varepsilon)} < \infty,$$

then

$$\rho(P_n, W) = O(n^{-1/8}(\log n)^{1/2})$$

as $n \rightarrow \infty$.

The absolutely regular condition is stronger than the strong mixing condition. Actually, for each n , $\alpha(n) \leq \beta(n)$. However (1.3) does not imply (1.2) in general, thus our theorem is not a complete generalization of Theorem A. Although it is true for some small $\varepsilon > 0$. But we emphasize that in our case we assume only the strong mixing condition. As far as the author knows, there are few results on the rate of convergence of the invariance principle under the strong mixing condition. Our technique of the proof is different from that of Theorem A.

2. Preliminaries

In this section, we state three lemmas. The first lemma is due to Davydov [4].

Lemma 1. Let p, q and u be positive numbers with $p^{-1} + q^{-1} + u^{-1} = 1$. Suppose that X is \mathcal{F}_1^a -measurable and Y is \mathcal{F}_{a+n}^∞ -measurable. Moreover suppose that $E|X|^p < \infty$ and $E|Y|^q < \infty$. Then

$$|E(XY) - E(X)E(Y)| \leq 10(\alpha(n))^{1/u}(E|X|^p)^{1/p}(E|Y|^q)^{1/q}.$$

Recently Yokoyama [9] gave an estimate of the p -th absolute moment of sums of strong mixing random variables as follows.

Lemma 2. Let $\{X_i\}$ be a strictly stationary strong mixing sequence with $EX_1 = 0$ and $E|X_1|^r < \infty$ for some $r > 2$. If there exists p with $2 < p < r$ such that

$$(2.1) \quad \sum_{i=1}^{\infty} i^{p/2-1}(\alpha(i))^{(r-p)/r} < \infty,$$

then there exists a positive constant K such that

$$E \left| \sum_{i=1}^n X_i \right|^p \leq Kn^{p/2}, \quad n \geq 1.$$

In what follows, as an absolute positive constant, we shall use a K which may be different in the different equations.

For some $0 < v < 1$, let $M = [n/[n^v]] + 1$, where $[\cdot]$ denotes the integer part of \cdot . For $j = 1, \dots, M-1$, define $I_j = \{(j-1)[n^v] + 1, (j-1)[n^v] + 2, \dots, j[n^v]\}$ and $I_M =$

$\{(M-1)[n^\nu]+1, (M-1)[n^\nu]+2, \dots, n\}$. Let

$$y_j = \sum_{i \in I_j} n^{-1/2} X_i \quad \text{for } j=1, \dots, M.$$

Using Lemmas 1 and 2 we can easily prove the following lemma.

Lemma 3. *Under the conditions of Theorem 1, for $j=1, \dots, M-1$,*

$$(2.2) \quad |Ey_j^2 - n^{\nu-1}| \leq Kn^{-1/2},$$

and for any p with $2 < p < 2rs/(r+s)$,

$$(2.3) \quad E|y_j|^p \leq Kn^{(\nu-1)p/2}.$$

Also we have

$$(2.4) \quad |Ey_M^2 - (n - (M-1)[n^\nu])n^{-1}| \leq Kn^{-1}$$

and

$$(2.5) \quad E|y_M|^p \leq Kn^{(\nu-1)p/2}.$$

Proof. From the stationarity, it suffices to show (2.2) only for $j=1$. By the ordinary argument (see e.g. Philipp-Stout [6], p. 28),

$$\begin{aligned} Ey_1^2 &= n^{-1} \left\{ \sum_{i=1}^{[n^\nu]} EX_i^2 + 2 \sum_{i=1}^{[n^\nu]-1} ([n^\nu] - i) E(X_1 X_{i+1}) \right\} \\ &= n^{-1} \left\{ [n^\nu] \left(EX_1^2 + 2 \sum_{i=1}^{\infty} E(X_1 X_{i+1}) \right) \right. \\ &\quad \left. - 2[n^\nu] \sum_{i=[n^\nu]}^{\infty} E(X_1 X_{i+1}) - 2 \sum_{i=1}^{[n^\nu]-1} i E(X_1 X_{i+1}) \right\}. \end{aligned}$$

Applying Lemma 1 we have

$$|E(X_1 X_{i+1})| \leq 10(\alpha(i))^{(r-2)/r} (E|X_1|^r)^{1/r} (E|X_{i+1}|^r)^{1/r}.$$

Thus from condition (1.2) and the monotonicity of $\{\alpha(i)\}$, it follows that

$$[n^\nu] \sum_{i=[n^\nu]}^{\infty} E(X_1 X_{i+1}) \leq K[n^\nu] \sum_{i=[n^\nu]}^{\infty} (\alpha(i))^{(r-2)/r} \leq K,$$

and

$$\sum_{i=1}^{[n^\nu]-1} i E(X_1 X_{i+1}) \leq K \sum_{i=1}^{[n^\nu]-1} i (\alpha(i))^{(r-2)/r} \leq K.$$

Hence we have (2.2) for $j=1$.

Furthermore, if $2 < p \leq 2rs/(r+s)$, then condition (1.2) implies

$$\sum_{i=1}^{\infty} i^{-1+p/2} (\alpha(i))^{(r-s)/r} < \infty,$$

from which (2.3) follows because of Lemma 2. (2.4) and (2.5) are similarly shown.

3. Martingale representations

For $i=1, \dots, M$, let \mathcal{L}_i be the σ -field generated by the random variables $\{y_1, \dots, y_i\}$ and $\mathcal{L}_0 = \{\Omega, \phi\}$. For simplicity let $\mathcal{F}_j = \mathcal{F}_1^j$ for $j=1, 2, \dots$ and $\mathcal{F}_0 = \{\Omega, \phi\}$.

The idea of the proofs of the following lemmas in this section is the same as in the proof of Lemma 7.4.1 in Philipp-Stout [6].

Lemma 4. *Under the assumptions of Theorem 1, we can represent X_i in the form,*

$$(3.1) \quad X_i = \eta_i - d_{i+1} + d_i, \quad i=1, \dots, n,$$

where $\{\eta_i\}$ is a martingale difference sequence with respect to $\{\mathcal{F}_i\}$ and where

$$(3.2) \quad \|d_i\|_s \leq K \quad \text{and} \quad \|\eta_i\|_s \leq K, \quad i=1, \dots, n,$$

where $\|\cdot\|_s = (E|\cdot|^s)^{1/s}$.

Proof. Define

$$d_i = \sum_{k=0}^{\infty} E(X_{i+k} | \mathcal{F}_{i-1})$$

and

$$\eta_i = X_i - d_i + d_{i+1}$$

for $1 \leq i \leq n$. Then the representation (3.1) follows.

On the other hand, by the Minkovsky inequality, we have

$$\|d_i\|_s \leq \sum_{k=0}^{\infty} \|E(X_{i+k} | \mathcal{F}_{i-1})\|_s.$$

Using Lemma 1 with $p=r$, $q=s/(s-1)$ and $u=sr/(r-s)$, and noting $EX_{i+k}=0$, we have

$$\begin{aligned} E|E(X_{i+k} | \mathcal{F}_{i-1})|^s &= E(E(X_{i+k} | \mathcal{F}_{i-1})E(X_{i+k} | \mathcal{F}_{i-1}) | E(X_{i+k} | \mathcal{F}_{i-1})|^{s-2}) \\ &= E(X_{i+k}E(X_{i+k} | \mathcal{F}_{i-1}) | E(X_{i+k} | \mathcal{F}_{i-1})|^{s-2}) \\ &\leq 10(\alpha(k+1))^{(r-s)/rs} \|X_{i+k}\|_r \| |E(X_{i+k} | \mathcal{F}_{i-1})|^{s-1} \|_{s/(s-1)}. \end{aligned}$$

Thus we have by assumption (1.2),

$$\|d_i\|_s \leq 10 \|X_1\|_r \sum_{k=0}^{\infty} (\alpha(k+1))^{(r-s)/rs} \leq K,$$

and also by the Minkovsky inequality,

$$\|\eta_i\|_s \leq \|X_i\|_s + \|d_i\|_s + \|d_{i+1}\|_s \leq K.$$

The lemma is concluded.

Lemma 5. *We can represent y_i in the form*

$$(3.3) \quad y_i = \xi_i + v_i, \quad i = 1, \dots, M,$$

where $\{\xi_i\}$ is a martingale difference sequence with respect to $\{\mathcal{L}_i\}$ and where

$$(3.4) \quad \|v_i\|_s \leq Kn^{-1/2} \quad \text{and} \quad \|\xi_i\|_s \leq Kn^{(v-1)/2}, \quad i = 1, \dots, M.$$

Proof. Let $\xi_i = y_i - E(y_i | \mathcal{L}_{i-1})$, and $v_i = E(y_i | \mathcal{L}_{i-1})$, then we have the representation (3.3). For the proof of (3.4), we note that

$$\begin{aligned} E |E(y_i | \mathcal{L}_{i-1})|^s &= \sum_{k=1}^{[n^v]} E \{ n^{-1/2} X_{(i-1)[n^v]+k} E(y_i | \mathcal{L}_{i-1}) | E(y_i | \mathcal{L}_{i-1}) |^{s-2} \} \\ &\leq \sum_{k=1}^{[n^v]} 10n^{-1/2} (\alpha(k))^{(r-s)/rs} \|X_{(i-1)[n^v]+k}\|_r \|E(y_i | \mathcal{L}_{i-1})\|_s^{s-1} \\ &\leq Kn^{-1/2} \|E(y_i | \mathcal{L}_{i-1})\|_s^{s-1}, \end{aligned}$$

by Lemma 1 and assumption (1.2). Then the proof of the lemma is concluded by the same reasoning as in the proof of Lemma 4.

Lemma 6. *We can represent v_i in the form*

$$(3.5) \quad v_i = \theta_i - g_{i+1} + g_i, \quad i = 1, \dots, M,$$

where $\{\theta_i\}$ is a martingale difference sequence with respect to $\{\mathcal{L}_{i-1}\}$ and where

$$(3.6) \quad \|g_i\|_s \leq Kn^{-1/2} \quad \text{and} \quad \|\theta_i\|_s \leq Kn^{-1/2}, \quad i = 1, \dots, M.$$

Proof. Define

$$g_i = \sum_{k=0}^{M-i} E(y_{i+k} | \mathcal{L}_{i-2}) \quad \text{and} \quad \theta_i = v_i - g_{i+1} + g_i \quad \text{for } 1 \leq i \leq M,$$

where $\mathcal{L}_{-1} = \mathcal{L}_0 = \{\phi, \Omega\}$. Then the representation (3.5) is obtained.

Using Lemma 1 we have

$$\begin{aligned} E |E(y_{i+k} | \mathcal{L}_{i-2})|^s &= \sum_{m=1}^{[n^v]} E \{ n^{-1/2} X_{(i+k-1)[n^v]+m} E(y_{i+k} | \mathcal{L}_{i-2}) | E(y_{i+k} | \mathcal{L}_{i-2}) |^{s-2} \} \\ &\leq \sum_{m=1}^{[n^v]} 10n^{-1/2} (\alpha((k+1)[n^v]+m))^{(r-s)/rs} \\ &\quad \times \|X_{(i+k-1)[n^v]+m}\|_r \|E(y_{i+k} | \mathcal{L}_{i-2})\|_s^{s-1}. \end{aligned}$$

Thus from the Minkovsky inequality and assumption (1.2) we have

$$\begin{aligned} \|g_i\|_s &\leq \sum_{k=0}^{M-i} \|E(y_{i+k} | \mathcal{L}_{i-2})\|_s \\ &\leq Kn^{-1/2} \sum_{k=0}^{M-i} \sum_{m=1}^{[n^v]} (\alpha((k+1)[n^v]+m))^{(r-s)/rs} \end{aligned}$$

$$\begin{aligned}
 &= Kn^{-1/2} \sum_{k=[n^v]+1}^{(M-i+2)[n^v]} (\alpha(k))^{(r-s)/rs} \\
 &\leq Kn^{-1/2}, \quad i=1, \dots, M.
 \end{aligned}$$

The rest of the proof is similar to that of Lemma 4.

Lemma 7. *We can represent $y_i v_i$ in the form*

$$(3.7) \quad y_i v_i = \mu_i + u_i, \quad i=1, \dots, M,$$

where $\{\mu_i\}$ is a martingale difference sequence with respect to $\{\mathcal{L}_i\}$ and where

$$(3.8) \quad \|u_i\|_{s/2} \leq Kn^{-1/2} \quad \text{and} \quad \|\mu_i\|_{s/2} \leq Kn^{(v-2)/2}, \quad i=1, \dots, M.$$

Proof. Let $u_i = E(y_i v_i | \mathcal{L}_{i-1})$ and $\mu_i = y_i v_i - u_i$ for $1 \leq i \leq M$. Then the representation (3.7) follows. Since

$$u_i = E(y_i E(y_i | \mathcal{L}_{i-1}) | \mathcal{L}_{i-1}) = (E(y_i | \mathcal{L}_{i-1}))^2 = v_i^2,$$

it follows from (3.4) in Lemma 5 that

$$\|u_i\|_{s/2} \leq \|v_i\|_s^2 \leq Kn^{-1}, \quad i=1, \dots, M.$$

On the other hand by the Schwartz inequality and Lemmas 3 and 5, we have

$$\|y_i v_i\|_{s/2} \leq \|y_i\|_s \|v_i\|_s \leq Kn^{(v-2)/2}.$$

Thus by the Minkovsky inequality we have

$$\|\mu_i\|_{s/2} \leq \|y_i v_i\|_{s/2} + \|u_i\|_{s/2} \leq Kn^{(v-2)/2}, \quad i=1, \dots, M,$$

and the statement of the lemma is concluded.

Lemma 8. *Let $h_i = y_i^2 - E y_i^2$. We can represent h_i in the form*

$$(3.9) \quad h_i = \zeta_i - w_{i+1} + w_i, \quad i=1, \dots, M,$$

where $\{\zeta_i\}$ is a martingale difference sequence with respect to $\{\mathcal{L}_i\}$ and where

$$(3.10) \quad \|w_i\|_{s/2} \leq Kn^{v-1} \quad \text{and} \quad \|\zeta_i\|_{s/2} \leq n^{v-1}, \quad i=1, \dots, M.$$

Proof. Define

$$w_i = \sum_{k=0}^{M-i} E(h_{i+k} | \mathcal{L}_{i-1}) \quad \text{and} \quad \zeta_i = h_i + w_{i+1} - w_i \quad \text{for} \quad 1 \leq i \leq M,$$

where $w_{M+1} \equiv 0$. Then the representation (3.9) follows.

Let \mathcal{L}_a^b be σ -field generated by $\{h_a, h_{a+1}, \dots, h_b\}$. Since $\{h_i\}$ satisfies the strong mixing condition such that for any $1 \leq a \leq M-1$ and $k \geq 1$,

$$\sup_{A \in \mathcal{L}_1^a, B \in \mathcal{L}_{a+k}^M} |P(AB) - P(A)P(B)| \leq \alpha((k-1)[n^v] + 1),$$

we have from Lemma 1,

$$\begin{aligned}
E|E(h_{i+k}|\mathcal{L}_{i-1})|^{s/2} &= E(h_{i+k}E(h_{i+k}|\mathcal{L}_{i-1})|E(h_{i+k}|\mathcal{L}_{i-1})|^{-2+s/2}) \\
&\leq 10(\alpha(k[n^v]+1))^{(r-s)/rs}\|h_{i+k}\|_{rs/(r+s)}\|E(h_{i+k}|\mathcal{L}_{i-1})\|_{s/(s-2)}^{(s-2)/2} \\
&\leq K(\alpha(k[n^v]+1))^{(r-s)/rs}\|y_{i+k}\|_{2rs/(r+s)}^2\|E(h_{i+k}|\mathcal{L}_{i-1})\|_{s/2}^{(s-2)/2}.
\end{aligned}$$

Thus from the Minkovsky inequality, Lemma 3 and assumption (1.2) it follows that

$$\begin{aligned}
\|w_i\|_{s/2} &\leq \sum_{k=0}^{M-i} \|E(h_{i+k}|\mathcal{L}_{i-1})\|_{s/2} \\
&\leq Kn^{v-1} \sum_{k=0}^{M-i} (\alpha(k[n^v]+1))^{(r-s)/rs} \\
&\leq Kn^{v-1}
\end{aligned}$$

and

$$\|\zeta_i\|_{s/2} \leq \|y_i\|_s^2 + Ey_i^2 + \|w_{i+1}\|_{s/2} + \|w_i\|_s \leq Kn^{v-1}, \quad i=1, \dots, M.$$

Hence the lemma is proved.

4. Proof of Theorem 1

We first prove the theorem in case $2 < s \leq 4$. Let $v = (s-2)/2(s-1)$, $\varepsilon_n = n^{-\delta}$ and $\lambda_n = n^{-2\delta}(\log n)^{-2}$. If we could construct a Brownian motion $\{B(t), 0 \leq t \leq 1\}$ on a probability space (Ω, \mathcal{F}, P) such that

$$(4.1) \quad P\{d(X_n(t), B(t)) \geq \varepsilon_n\} = O(\varepsilon_n) = O(n^{-\delta})$$

as $n \rightarrow \infty$, then we have by the same reasoning as in the proof of Theorem 1 in [2] that

$$(4.2) \quad \rho(P_n, W) = O(n^{-\delta}) \quad \text{for any } \delta < s(s-2)/4(s-1)(s+1),$$

which concludes the theorem. Hence we first have to define a Brownian motion with the property (4.1). To do that, we use the following lemma due to Strassen [8].

Lemma 9. *Let $\{z_i\}$ be random variables such that for all $i \geq 1$, $E(z_i^2 | z_{i-1}, \dots, z_1)$ exists and $E(z_i | z_{i-1}, \dots, z_1) = 0$ a.s. Then on a probability space (Ω, \mathcal{F}, P) , there is a standard Brownian motion $\{B(t)\}$ together with a sequence of non-negative random variables $\{T_i\}$ such that the joint distributions of $\{z_i\}$ are the same as those of*

$$\left\{ B\left(\sum_{k=1}^i T_k\right) - B\left(\sum_{k=1}^{i-1} T_k\right) \right\}$$

Moreover, let \mathcal{B}_i be the σ -field generated by

$$\left\{ \{z_1, \dots, z_i\}; B(t), 0 \leq t \leq \sum_{k=1}^i T_k \right\}.$$

Then we see that for all $i \geq 1$, T_i is \mathcal{B}_i -measurable, $E(T_i | \mathcal{B}_{i-1})$ exists and

$$(4.3) \quad E(T_i | \mathcal{B}_{i-1}) = E(z_i^2 | \mathcal{B}_{i-1}) = E(z_i^2 | z_{i-1}, \dots, z_1) \quad \text{a.s.}$$

Furthermore, for all $i \geq 1$, if $E(|z_i|^r | z_{i-1}, \dots, z_1)$ exists for some $r > 2$, then

$$(4.4) \quad ET_i^{r/2} \leq KE|z_i|^r.$$

From Lemma 9 we can construct a Brownian motion $\{B^*(t)\}$ and a sequence of positive random variables $\{T_i^*\}$ on a probability space $(\Omega^*, \mathcal{F}^*, P^*)$ such that the joint distributions of $\{\xi_1, \xi_2, \dots, \xi_M\}$ are the same as those of

$$\left\{ B^*(T_1^*), B^*(T_1^* + T_2^*) - B^*(T_1^*), \dots, B^*\left(\sum_{i=1}^M T_i^*\right) - B^*\left(\sum_{i=1}^{M-1} T_i^*\right) \right\}.$$

Moreover using Lemma A1 in [1] we can redefine $(\{y_i\}, \{\xi_i\}, \{T_i\}, \{\hat{B}(t)\})$ on a common probability space (Ω, \mathcal{F}, P) such that the distributions of $(\{y_i\}, \{\xi_i\})$ and $(\{\hat{\xi}_i\}, \{\hat{T}_i\}, \{\hat{B}(t)\})$ are the same as those of $(\{y_i\}, \{\xi_i\})$ and

$$\left(\left\{ B^*\left(\sum_{k=1}^i T_k^*\right) - B^*\left(\sum_{k=1}^{i-1} T_k^*\right) \right\}, \{T_i^*\}, \{B^*(t)\} \right),$$

respectively.

In what follows, for simplicity, we shall write $\{y_i\}, \{\xi_i\}, \{T_i\}$, and $\{B(t)\}$ for $\{y_i\}, \{\hat{\xi}_i\}, \{\hat{T}_i\}$ and $\{\hat{B}(t)\}$.

Define M points on $[0, 1]$ such that $a_k = k[n^\nu]/n, k = 0, 1, \dots, M-1$ and $a_M = 1$. Let $\tilde{X}_n(t)$ and $\tilde{B}_n(t)$ be continuous polygonal lines defined by

$$\begin{aligned} \tilde{X}_n(t) &= X_n(a_k) + (X_n(a_{k+1}) - X_n(a_k))(t - a_k)/(a_{k+1} - a_k), \\ &\text{for } t \in [a_k, a_{k+1}], \quad k = 0, 1, \dots, M-1, \end{aligned}$$

and

$$\begin{aligned} \tilde{B}_n(t) &= B(a_k) + (B(a_{k+1}) - B(a_k))(t - a_k)/(a_{k+1} - a_k), \\ &\text{for } t \in [a_k, a_{k+1}], \quad k = 0, 1, \dots, M-1, \end{aligned}$$

respectively.

Lemma 10. As $n \rightarrow \infty$

$$P\{d(X_n(t), \tilde{X}_n(t)) \geq \varepsilon_n\} = o(\varepsilon_n).$$

Proof. From the definition of $\{X_n(t)\}$ and $\{\tilde{X}_n(t)\}$ we have

$$\begin{aligned} (4.5) \quad P\{d(X_n(t), \tilde{X}_n(t)) \geq \varepsilon_n\} &\leq MP \left\{ \max_{1 \leq k \leq [n^\nu]} \left| \sum_{i=1}^k n^{-1/2} X_i - (k/[n^\nu]) \sum_{i=1}^k n^{-1/2} X_i \right| \geq \varepsilon_n \right\} \\ &\leq MP \left\{ \max_{1 \leq k \leq [n^\nu]} \left| \sum_{i=1}^k n^{-1/2} X_i \right| \geq \varepsilon_n/2 \right\} \\ &\equiv A_1, \quad \text{say.} \end{aligned}$$

Recall that $M = [n/[n^\nu]] + 1 = O(n^{1-\nu})$. Using Lemma 4 we have

$$(4.6) \quad A_1 \leq M P \left\{ \max_{1 \leq k \leq [n^\nu]} \left| \sum_{i=1}^k n^{-1/2} \eta_i \right| \geq \varepsilon_n/4 \right\} \\ + M P \left\{ \max_{1 \leq k \leq [n^\nu]} \left| \sum_{i=1}^k n^{-1/2} (d_{i+1} - d_i) \right| \geq \varepsilon_n/4 \right\}.$$

Applying the Kolmogorov maximal inequality, the Burkholder inequality for martingale difference sequences (see e.g. [3]) and Lemma 4 to the first term on the right hand side of (4.6), and noting that $E|\eta_i|^s < \infty$, we have

$$(4.7) \quad M P \left\{ \max_{1 \leq k \leq [n^\nu]} \left| \sum_{i=1}^k n^{-1/2} \eta_i \right| \geq \varepsilon_n/4 \right\} \leq K M \varepsilon_n^{-s} E \left| \sum_{i=1}^{[n^\nu]} n^{-1/2} \eta_i \right|^s \\ \leq K M \varepsilon_n^{-s} n^{\nu(s-2)/2} \sum_{i=1}^{[n^\nu]} E |n^{-1/2} \eta_i|^s \\ \leq K \varepsilon_n^{-s} n^{-s(s-2)/4(s-1)} \\ = o(\varepsilon_n).$$

Moreover we have from the Chebyshev inequality and Lemma 4,

$$(4.8) \quad M P \left\{ \max_{1 \leq k \leq [n^\nu]} \left| \sum_{i=1}^k n^{-1/2} (d_{i+1} - d_i) \right| \geq \varepsilon_n/4 \right\} = M P \left\{ \max_{1 \leq k \leq [n^\nu]} |d_{k+1} - d_1| \geq n^{1/2} \varepsilon_n/4 \right\} \\ \leq M \sum_{k=1}^{[n^\nu]} P \{ |d_{k+1} - d_1| \geq n^{1/2} \varepsilon_n/4 \} \\ \leq M \sum_{k=1}^{[n^\nu]} (n^{1/2} \varepsilon_n/4)^{-s} E |d_{k+1} - d_1|^s \\ \leq K n^{1-s/2} \varepsilon_n^{-s} \\ = o(\varepsilon_n).$$

Thus we have $A_1 = o(\varepsilon_n)$ from (4.6), (4.7) and (4.8) and conclude this lemma. The following lemma is due to Borovkov (Lemma 2 in [2]).

Lemma 11. As $n \rightarrow \infty$,

$$P\{d(B(t), \tilde{B}_n(t)) \geq \varepsilon_n\} = o(\exp(-n^\delta)) = o(\varepsilon_n),$$

for some $\delta > 0$.

Finally we shall prove the following

Lemma 12. As $n \rightarrow \infty$,

$$P\{d(\tilde{X}_n(t), \tilde{B}_n(t)) \geq \varepsilon_n\} = o(\varepsilon_n).$$

Proof. By the definition of $\{\tilde{X}_n(t)\}$ and $\{\tilde{B}_n(t)\}$ and by Lemma 4 we have

(4.9)

$$\begin{aligned} P\{d(\tilde{X}_n(t), \tilde{B}_n(t)) \geq \varepsilon_n\} &= P\left\{ \max_{1 \leq k \leq M} \left| \sum_{i=1}^k y_i - B(a_k) \right| \geq \varepsilon_n \right\} \\ &\leq P\left\{ \max_{1 \leq k \leq M} \left| \sum_{i=1}^k \xi_i - B(a_k) \right| \geq \varepsilon_n/2 \right\} + P\left\{ \max_{1 \leq k \leq M} \left| \sum_{i=1}^k v_i \right| \geq \varepsilon_n/2 \right\} \\ &\equiv D_1 + D_2, \quad \text{say.} \end{aligned}$$

We first estimate D_1 . Denote

$$Z_k = \sum_{i=1}^k (T_i - ET_i).$$

By Lemma 5 we have

$$\begin{aligned} (4.10) \quad D_1 &\leq P\left\{ \max_{1 \leq k \leq M} \left| B\left(Z_k + \sum_{i=1}^k ET_i\right) - B(a_k) \right| \geq \varepsilon_n, \max_{1 \leq k \leq M} |Z_k| < 3\lambda_n \right\} \\ &\quad + P\left\{ \max_{1 \leq k \leq M} |Z_k| \geq 3\lambda_n \right\} \\ &\equiv E_1 + E_2, \quad \text{say.} \end{aligned}$$

According to the proof of Philipp-Stout [6], we use (4.3) in Lemma 9 to estimate E_2 . Then we have

$$\begin{aligned} (4.11) \quad E_2 &\leq P\left\{ \max_{1 \leq k \leq M} \left| \sum_{i=1}^k (T_i - E(T_i | \mathcal{B}_{i-1})) \right| \geq \lambda_n \right\} \\ &\quad + P\left\{ \max_{1 \leq k \leq M} \left| \sum_{i=1}^k (E(\xi_i^2 | \xi_{i-1}, \dots, \xi_1) - \xi_i^2) \right| \geq \lambda_n \right\} \\ &\quad + P\left\{ \max_{1 \leq k \leq M} \left| \sum_{i=1}^k (\xi_i^2 - E\xi_i^2) \right| \geq \lambda_n \right\} \\ &\equiv L_1 + L_2 + L_3, \quad \text{say.} \end{aligned}$$

Since $\{T_i - E(T_i | \mathcal{B}_{i-1})\}$ and $\{\xi_i^2 - E(\xi_i^2 | \xi_{i-1}, \dots, \xi_1)\}$ are martingale difference sequences, respectively, we have, using the Kolmogorov maximal inequality and the Burkholder inequality for martingale difference sequences,

$$L_1 \leq K\lambda_n^{-s/2} \sum_{i=1}^M E |T_i - E(T_i | \mathcal{B}_{i-1})|^{s/2}$$

and

$$L_2 \leq K\lambda_n^{-s/2} \sum_{i=1}^M E |\xi_i^2 - E(\xi_i^2 | \xi_{i-1}, \dots, \xi_1)|^{s/2}.$$

Using the Hölder inequality for the conditional expectation and Lemmas 9 and 5, we have

$$E|E(T_i|\mathcal{B}_{i-1})|^{s/2} \leq ET_i^{s/2} \leq KE|\xi_i|^s \leq Kn^{(v-1)s/2},$$

and

$$E|E(\xi_i^2|\xi_{i-1}, \dots, \xi_1)|^{s/2} \leq E|\xi_i|^s \leq Kn^{(v-1)s/2}.$$

Hence

$$(4.12) \quad L_1 \leq Kn^{(v-1)(s-2)/2} \lambda_n^{-s/2} = o(\varepsilon_n)$$

and

$$(4.13) \quad L_2 = o(\varepsilon_n).$$

We next estimate L_3 . By the definition of ξ_i we have

$$(4.14) \quad L_3 \leq P \left\{ \max_{1 \leq k \leq M} \left| \sum_{i=1}^k (y_i^2 - Ey_i^2) \right| \geq \lambda_n/3 \right\} + P \left\{ \max_{1 \leq k \leq M} \left| \sum_{i=1}^k (y_i v_i - E(y_i v_i)) \right| \geq \lambda_n/6 \right\} \\ + P \left\{ \max_{1 \leq k \leq M} \left| \sum_{i=1}^k (v_i^2 - Ev_i^2) \right| \geq \lambda_n/3 \right\} \\ \equiv L_{31} + L_{32} + L_{33}, \quad \text{say.}$$

As to L_{31} , using Lemma 8, the Kolmogorov maximal inequality and the Burkholder inequality for martingale difference sequences, we have

$$(4.15) \quad L_{31} \leq P \left\{ \max_{1 \leq k \leq M} \left| \sum_{i=1}^k \zeta_i \right| \geq \lambda_n/6 \right\} + P \left\{ \max_{1 \leq k \leq M} \left| \sum_{i=1}^k (w_{i+1} - w_1) \right| \geq \lambda_n/6 \right\} \\ \leq K\lambda_n^{-s/2} E \left| \sum_{i=1}^M \zeta_i \right|^{s/2} + \sum_{i=1}^M P\{|w_{i+1} - w_1| \geq \lambda_n/6\} \\ \leq K\lambda_n^{-s/2} \sum_{i=1}^M E|\zeta_i|^{s/2} + K\lambda_n^{-s/2} \sum_{i=1}^M E|w_{i+1} - w_1|^s \\ \leq Kn^{(v-1)(s-2)/2} \lambda_n^{-s/2} \\ = o(\varepsilon_n).$$

We next handle L_{32} . Note that from Lemmas 5 and 3,

$$\sum_{i=1}^M |E(y_i v_i)| = \sum_{i=1}^M |E(y_i(Y_i | \mathcal{L}_{i-1}))| = \sum_{i=1}^M Ey_i^2 = O(n^{-v}) = o(\lambda_n).$$

Thus from Lemma 7 we have

$$(4.16) \quad L_{32} \leq P \left\{ \max_{1 \leq k \leq M} \left| \sum_{i=1}^k y_i v_i \right| \geq \lambda_n/6 - \sum_{i=1}^M |E(y_i v_i)| \right\} \\ \leq P \left\{ \max_{1 \leq k \leq M} \left| \sum_{i=1}^k y_i v_i \right| \geq \lambda_n/8 \right\}$$

$$\leq P \left\{ \max_{1 \leq k \leq M} \left| \sum_{i=1}^k \mu_i \right| \geq \lambda_n/16 \right\} + P \left\{ \max_{1 \leq k \leq M} \left| \sum_{i=1}^k u_i \right| \geq \lambda_n/16 \right\}.$$

From the Chebyshev inequality, the Minkovsky inequality and Lemma 7, it follows that

$$(4.17) \quad \begin{aligned} P \left\{ \max_{1 \leq k \leq M} \left| \sum_{i=1}^k u_i \right| \geq \lambda_n/16 \right\} &\leq K \lambda_n^{-s/2} E \left(\sum_{i=1}^M |u_i| \right)^{s/2} \\ &\leq K \lambda_n^{-s/2} \left(\sum_{i=1}^M \|u_i\|_{s/2} \right)^{s/2} \\ &\leq K \lambda_n^{-s/2} n^{(1-2\nu)s/4} \\ &= o(\varepsilon_n). \end{aligned}$$

Since $\{\mu_i\}$ is a martingale difference sequence, we have

$$(4.18) \quad \begin{aligned} P \left\{ \max_{1 \leq k \leq M} \left| \sum_{i=1}^k \mu_i \right| \geq \lambda_n/16 \right\} &\leq K \lambda_n^{-s/2} E \left| \sum_{i=1}^M \mu_i \right|^{s/2} \\ &\leq K \lambda_n^{-s/2} \sum_{i=1}^M E |\mu_i|^{s/2} \\ &\leq K \lambda_n^{-s/2} \sum_{i=1}^M (E |y_i v_i|^{s/2} + E |u_i|^{s/2}). \end{aligned}$$

From Lemma 7 we have

$$E |y_i v_i|^{s/2} \leq K n^{(\nu-2)s/4} \quad \text{and} \quad E |u_i|^{s/2} \leq K n^{-s/2},$$

thus it follows that

$$(4.18) \quad \begin{aligned} P \left\{ \max_{1 \leq k \leq M} \left| \sum_{i=1}^k \mu_i \right| \geq \lambda_n/16 \right\} &\leq K \lambda_n^{-s/2} n^{1-s/2+(s-4)\nu/4} + K \lambda_n^{-s/2} n^{1-\nu-s/2} \\ &= o(\varepsilon_n). \end{aligned}$$

Combining (4.16), (4.17) and (4.18) we have

$$(4.19) \quad L_{32} = o(\varepsilon_n).$$

As to L_{33} , we have from Lemma 5

$$(4.20) \quad \begin{aligned} L_{33} &\leq P \left\{ \max_{1 \leq k \leq M} \sum_{i=1}^k v_i^2 \geq \lambda_n/3 - \sum_{i=1}^M E v_i^2 \right\} \\ &\leq P \left\{ \sum_{i=1}^M v_i^2 \geq \lambda_n/3 - o(\lambda_n) \right\} \\ &\leq P \left\{ \sum_{i=1}^M v_i^2 \geq \lambda_n/4 \right\} \end{aligned}$$

$$\begin{aligned}
&\leq K\lambda_n^{-s/2} E\left(\sum_{i=1}^M v_i^2\right)^{s/2} \\
&\leq K\lambda_n^{-s/2} \left(\sum_{i=1}^M \|v_i^2\|_{s/2}\right)^{s/2} \\
&\leq Kn^{-vs/2} \lambda_n^{-s/2} \\
&= o(\varepsilon_n),
\end{aligned}$$

From (4.14), (4.15), (4.19) and (4.20), it follows that as $n \rightarrow \infty$,

$$(4.21) \quad L_3 = o(\varepsilon_n).$$

From (4.11), (4.12), (4.13) and (4.21) we obtain

$$(4.22) \quad E_2 = o(\varepsilon_n).$$

We next estimate E_1 in (4.10) following Rosenkrantz [7]. Since $E(y_i v_i) = E v_i^2$ for all $1 \leq i \leq M$, we have from Lemmas 3, 5 and 9

$$(4.23) \quad \max_{1 \leq k \leq M} \left| \sum_{i=1}^k E T_i - a_k \right| \leq \max_{1 \leq k \leq M} \left| \sum_{i=1}^k E y_i^2 - a_k \right| + \sum_{i=1}^M E v_i^2 = o(\lambda_n).$$

Recalling

$$Z_k = \sum_{i=1}^k (T_i - E T_i),$$

we have from (4.23)

$$\begin{aligned}
(4.24) \quad E_1 &= P \left\{ \max_{1 \leq k \leq M} \left| B\left(Z_k + \sum_{i=1}^k E T_i\right) - B(a_k) \right| \geq \varepsilon_n/2, \max_{1 \leq j \leq M} |Z_j| < 3\lambda_n \right\} \\
&\leq \sum_{k=1}^M P \left\{ \sup_{|t| \leq \max_{1 \leq j \leq M} |Z_j|} \left| \sum_{i=1}^j E T_i - a_k \right| + 3\lambda_n \left| B(t + a_k) - B(a_k) \right| \geq \varepsilon_n/2 \right\} \\
&\leq \sum_{k=1}^M P \left\{ \sup_{|t| \leq 4\lambda_n} |B(t + a_k) - B(a_k)| \geq \varepsilon_n/2 \right\} \\
&\leq 2MP \left\{ \sup_{0 \leq t \leq 4\lambda_n} |B(t)| \geq \varepsilon_n/2 \right\} \\
&\leq 8MP \{ |B(t)| \geq \varepsilon_n/(4\lambda_n^{1/2}) \} \\
&= o(\varepsilon_n).
\end{aligned}$$

The relations (4.9), (4.22) and (4.24) implies

$$(4.25) \quad D_1 = o(\varepsilon_n).$$

We shall finally estimate D_2 in (4.9). By Lemma 6 we have

$$(4.26) \quad D_2 \leq P \left\{ \max_{1 \leq k \leq M} \left| \sum_{i=1}^k \theta_i \right| \geq \varepsilon_n/4 \right\} + P \left\{ \max_{1 \leq k \leq M} \left| \sum_{i=1}^k (g_i - g_{i+1}) \right| \geq \varepsilon_n/4 \right\}$$

We first estimate the second term on the right hand side of (4.26) by Lemma 6 as follows.

$$(4.27) \quad \begin{aligned} P \left\{ \max_{1 \leq k \leq M} \left| \sum_{i=1}^k (g_i - g_{i+1}) \right| \geq \varepsilon_n \right\} &\leq P \left\{ \max_{1 \leq k \leq M} |g_1 - g_{k+1}| \geq \varepsilon_n \right\} \\ &\leq \sum_{k=1}^M P \{ |g_1 - g_{k+1}| \geq \varepsilon_n \} \\ &\leq \sum_{k=1}^M \varepsilon_n^{-s} E |g_1 - g_{k+1}|^s \\ &= o(\varepsilon_n). \end{aligned}$$

On the other hand, since $\{\theta_i\}$ is a martingale difference sequence, using the Kolmogorov maximal inequality, the Burkholder inequality for martingale difference sequence and Lemma 6, we have

$$(4.28) \quad \begin{aligned} P \left\{ \max_{1 \leq k \leq M} \left| \sum_{i=1}^k \theta_i \right| \geq \varepsilon_n \right\} &\leq KP \left\{ \left| \sum_{i=1}^M \theta_i \right| \geq \varepsilon_n \right\} \\ &\leq K \varepsilon_n^{-s} E \left| \sum_{i=1}^M \theta_i \right|^s \\ &\leq K \varepsilon_n^{-s} M^{(s-2)/2} \sum_{i=1}^M E |\theta_i|^s \\ &\leq K n^{-\nu s/2} \varepsilon_n^{-s} \\ &= o(\varepsilon_n). \end{aligned}$$

From (4.26), (4.27) and (4.28) we have

$$(4.29) \quad D_2 = o(\varepsilon_n).$$

The proof of Lemma 12 is completed from (4.9), (4.25) and (4.29).

(4.1) follows from Lemmas 10, 11 and 12 and the first part of Theorem 1 is concluded.

We next prove the second part of Theorem 1, where $r > 4$ and (1.2) holds for some s with $4 < s < r$. Let $\nu = 1/3$, $\varepsilon_n = n^{-2/15 - \kappa}$ for any $\kappa < (s-4)/30(1+s)$ and $\lambda_n = n^{-4/15 - 2\kappa} (\log n)^{-2}$. Then, by the Burkholder inequality, we have

$$\begin{aligned} L_1 &\leq K \lambda_n^{-s/2} M^{(s-4)/4} \sum_{i=1}^M E |T_i - E(T_i | \mathcal{B}_{i-1})|^{s/2} \\ &\leq K n^{(\nu-1)s/4} \lambda_n^{-s/2} \\ &= o(\varepsilon_n) \end{aligned}$$

and

$$\begin{aligned} L_2 &\leq K\lambda_n^{-s/2} M^{(s-4)/4} \sum_{i=1}^M E |\xi_i^2 - E(\xi_i^2 | \xi_{i-1}, \dots, \xi_1)|^{s/2} \\ &= o(\varepsilon_n). \end{aligned}$$

On the other hand, from the Burkholder inequality, it also follows that

$$\begin{aligned} P \left\{ \max_{1 \leq k \leq M} \left| \sum_{i=1}^k \zeta_i \right| \geq \lambda_n/6 \right\} &\leq K\lambda_n^{-s/2} E \left| \sum_{i=1}^M \zeta_i \right|^{s/2} \\ &\leq K\lambda_n^{-s/2} M^{(s-4)/4} \sum_{i=1}^M E |\zeta_i|^{s/2} \\ &\leq Kn^{(v-1)s/4} \lambda_n^{-s/2} \\ &= o(\varepsilon_n) \end{aligned}$$

and

$$\begin{aligned} P \left\{ \max_{1 \leq k \leq M} \left| \sum_{i=1}^k \mu_i \right| \geq \lambda_n \right\} &\leq K\lambda_n^{-s/2} E \left| \sum_{i=1}^M \mu_i \right|^{s/2} \\ &\leq K\lambda_n^{-s/2} M^{(s-4)/4} \sum_{i=1}^M E |\mu_i|^{s/2} \\ &\leq Kn^{-s/4} \lambda_n^{-s/2} + Kn^{-(1+v)s/4} \lambda_n^{-s/2} \\ &= o(\varepsilon_n). \end{aligned}$$

Thus we have $L_3 = o(\varepsilon_n)$ so that $E_2 = o(\varepsilon_n)$ by (4.11).

The rest of the proof is the same as that of the first part of Theorem 1, and we have

$$\rho(P_n, W) = o(\varepsilon_n) = o(n^{-2/15 - \kappa}).$$

The theorem is thus completely proved.

5. Proof of Corollary

If

$$\alpha(n) = O(e^{-\gamma n}),$$

for some $\gamma > 0$, then (1.2) holds for any $s < r$. Hence the conclusion of Theorem 1 holds for any

$$\delta < r(r-2)/4(r-1)(r+1),$$

in case $r \leq 4$. The case $r > 4$ is also handled in the same way.

Acknowledgement. I wish to thank Professor T. Kawata and Professor M. Maejima of Keio University for many detailed and helpful comments on this paper.

References

- [1] Berkes, I. and Philipp, W.: *Approximation theorems for independent and weakly dependent random variables*, Ann. Prob., **7** (1979), 29–54.
- [2] Borovkov, A. A.: *On the rate of convergence for the invariance principle*, Theor. Prob. Appl., **18** (1973), 207–225.
- [3] Burkholder, D. L.: *Martingale transforms*, Ann. Math. Statist., **37** (1966), 1497–1504.
- [4] Davydov, Yu. A.: *The invariance principle for stationary processes*, Theor. Prob. Appl., **15** (1970), 487–498.
- [5] Ibragimov, I. A. and Linnik, Yu. V.: *Independent and Stationary Sequences of Random Variables*, Wolters-Noordhoff, Groningen (1971).
- [6] Philipp, W. and Stout, W.: *Almost sure invariance principles for partial sums of weakly dependent random variables*, Mem. Amer. Math. Soc., **161** (1975).
- [7] Rosenkrantz, W. A.: *On rates of convergence for the invariance principle*, Trans. Amer. Math. Soc., **192** (1967), 542–552.
- [8] Strassen, V.: *Almost sure behaviour of sums of independent random variables and martingales*, Proc. Fifth Berkeley Symp. Math. Statist. Prob., **2** (1965), 315–343.
- [9] Yokoyama, R.: *Moments bounds for stationary mixing sequences*, Z. Wahrsch. Verw. Gebiete., **52** (1980), 45–57.
- [10] Yoshihara, K.: *Convergence rates of the invariance principle for absolutely regular sequences*, Yokohama Math. J., **27** (1979), 49–55.

Department of Mathematics
Keio University
Hiyoshi, Kohoku-ku
Yokohama 223, Japan