

A RANDOM RENEWAL THEOREM

By

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1. Introduction. Let $\{\xi_n\}$ be a sequence of i.i.d. random probability measures on the real line R with the expectation measure $F = E\xi_1$ and the second moment measure $G = E(\xi_1 \times \xi_1)$. The random renewal measure ζ for $\{\xi_n\}$ is defined by

$$\zeta = \delta_0 + \sum_{n=1}^{\infty} \xi_1 * \cdots * \xi_n .$$

The purpose of this note is to show that ζ satisfies an analogue of Blackwell type renewal theorem under some natural assumptions on F and G .

In §2 it is shown that ζ is locally finite a.s. iff F generates a transient random walk. In §3 we prove the following random renewal theorem: if F generates a transient random walk and if G is not supported by the diagonal of R^2 then $\lim_{t \rightarrow \infty} \zeta(I+t) = c \cdot \lambda(I)$ in probability, where I is a bounded interval of length $\lambda(I)$ and c is a constant. When ξ_n 's are non-random this result reduces to the well-known classical renewal theorem. Whether the above result holds in the sense of almost sure convergence or not is an open question.

2. Random renewal measures. Let M denote the set of all locally finite nonnegative measures ϕ on (R, \mathcal{B}) , where \mathcal{B} is the Borel σ -algebra of R . The smallest σ -algebra of subsets of M which respect to which every mapping $\phi \rightarrow \phi(A)$, $A \in \mathcal{B}$, is measurable is denoted by \mathcal{M} . The subset of M consisting of all probability measures is denoted by M_1 . Obviously $M_1 = \{\phi; \phi(R) = 1\} \in \mathcal{M}$. An M -valued random element ζ defined on a probability space (Ω, \mathcal{F}, P) is called a random measure (r.m.). A r.m. ξ satisfying $P\{\xi \in M_1\} = 1$ is called a random probability measure (r.p.m.). The expectation measure $E\zeta$ of a r.m. ζ is a (not necessarily locally finite) measure defined by $(E\zeta)(A) = E[\zeta(A)]$, $A \in \mathcal{B}$. If ξ is a r.p.m. then $E\xi \in M_1$. In this case write F_ξ for $E\xi$. We have

$$E \int f(t) \zeta(dt) = \int f(t) (E\zeta)(dt)$$

for every Baire function $f \geq 0$, where an integral sign without the limits means

integration over R . Thus if f is integrable with respect to $E\zeta$ then $\int f d\zeta$ is finite a.s. The second moment measure $E(\zeta \times \zeta)$ of a r.m. ζ is a (not necessarily locally finite) measure on $(R, \mathcal{R}) \times (R, \mathcal{R})$ defined by

$$[E(\zeta \times \zeta)](A \times A) = E[\zeta(A) \cdot \zeta(A)], \quad A, B \in \mathcal{R}.$$

If ξ is a r.p.m. then $G_\xi \equiv E(\xi \times \xi)$ is a probability measure.

The convolution $\xi * \eta$ of two r.m.'s ξ and η is not a r.m. in general. However if ξ and η are independent with $E\xi \in M$, $E\eta \in M$ and $(E\xi) * (E\eta) \in M$ then $\xi * \eta$ is a r.m. satisfying $E(\xi * \eta) = (E\xi) * (E\eta)$. In fact

$$\begin{aligned} E[\xi * \eta(A)] &= E \iint \chi_A(s+t) \xi(ds) \eta(dt) \\ &= \iint \chi_A(s+t) (E\xi)(ds) (E\eta)(dt) \\ &= (E\xi) * (E\eta)(A), \quad A \in \mathcal{R}, \end{aligned}$$

where χ_A is the indicator of A . In particular if ξ and η are independent r.p.m.'s then $\xi * \eta$ is a r.p.m. and we have $F_{\xi * \eta} = F_\xi * F_\eta$, $G_{\xi * \eta} = G_\xi * G_\eta$. The second equality follows from

$$\begin{aligned} E[\xi * \eta(A) \cdot \xi * \eta(B)] &= E \left[\int \xi(A-s) \eta(ds) \cdot \int \xi(B-t) \eta(dt) \right] \\ &= E \left[\int (\xi \times \xi)((A-s) \times (B-t)) (\eta \times \eta)(ds dt) \right] \\ &= E[(\xi \times \xi) * (\eta \times \eta)(A \times B)], \quad A, B \in \mathcal{R}. \end{aligned}$$

Let $\{\xi_n\}$ be a sequence of i.i.d.r.p.m.'s with common distribution P . Let $F = E\xi_1$, $G = E(\xi_1 \times \xi_1)$ and $U = \sum_{n=0}^{\infty} F^{n*}$, where $F^{0*} = \delta_0$ is the Dirac measure located at 0. We call F transient (recurrent) if F generates a transient (recurrent) random walk. Define the random renewal measure ζ by

$$\zeta = \delta_0 + \sum_{n=1}^{\infty} \xi_1 * \cdots * \xi_n.$$

Then we have

$$E\zeta = \delta_0 + \sum_{n=1}^{\infty} E(\xi_1 * \cdots * \xi_n) = \delta_0 + \sum_{n=1}^{\infty} F^{n*} = U.$$

The following theorem is an extension of a classical result on the recurrence of random walks due to Chung and Fuchs [3]. In particular if $\xi_n = \delta_{X_n}$, where $\{X_n\}$ is a sequence of i.i.d. random variables and δ_a is the Dirac measure at a , then ζ reduces to the ordinary renewal process: $\zeta = \sum_{n=0}^{\infty} \delta_{S_n}$, $S_n = \sum_{i=1}^n X_i$, $S_0 = 0$, and the

theorem reduces to the well-known result [3].

Theorem 1. *If F is recurrent then $\xi(I)=\infty$ a.s. for every neighborhood I of 0. If F is transient then $\zeta(I)<\infty$ a.s. for every bounded interval I .*

Proof. The second half is immediate. In fact if F is transient then $E\zeta(I)=U(I)<\infty$ and therefore $\zeta(I)<\infty$ a.s. for bounded I . To show the first half let $\{X_n\}$ be a sequence of i.i.d. random variables uniformly distributed over $[0, 1]$ and independent of $\{\xi_n\}$. Define a sequence $\{Y_n\}$ of random variables by $Y_n = \sup\{y; \xi_n(-\infty, y] \leq X_n\}$. Then we have easily

$$P\{Y_n \leq y\} = E[\xi_n(-\infty, y)] = F(-\infty, y).$$

Thus $\{Y_n\}$ is a sequence of i.i.d. random variables with common distribution F . Let $T_n = \sum_{k=0}^n Y_k$, $T_0=0$. It is easy to see that

$$P\{T_n \in A | \xi_1, \xi_2, \dots\} = \xi_1 * \dots * \xi_n(A) \text{ a.s.}$$

and therefore

$$\zeta(A) = P\left\{ \sum_{n=0}^{\infty} \delta_{T_n}(A) | \xi_1, \xi_2, \dots \right\} \text{ a.s.}$$

for every $A \in \mathcal{P}$. If F is recurrent then $\sum_{n=0}^{\infty} \delta_{T_n}(I) = \infty$ a.s. and therefore $\zeta(I) = \infty$ a.s. for every neighborhood I of 0.

3. A random renewal theorem. Throughout the rest suppose F is transient. Write μ for $\int tF(dt)$ if this integral exists. Let $D = \{(x, x); x \in R\}$ be the diagonal of R^2 and let M_d denote the set of all degenerate probability measures on R . Then the second moment measure G of a r.p.m. ξ is supported by D iff $P\{\xi \in M_d\} = 1$.

Theorem 2. *Suppose that F is transient non-arithmetic. If $G = G_{\xi_1}$ is not supported by D then*

$$(1) \quad \lim_{t \rightarrow \infty} \zeta(I+t) = c \cdot \lambda(I) \text{ in probability}$$

for every bounded interval I , where λ is the Lebesgue measure, $c = \mu^{-1}$ if $0 < \mu < \infty$ and $c = 0$ otherwise. When $c > 0$ the above condition on G is necessary for (1).

Proof. The last statement is obvious since if $0 < \mu < \infty$ and if G is supported by D then ξ_n is written as $\xi_n = \delta_{X_n}$ where $\{X_n\}$ is a sequence of i.i.d. random variables. In this case it is obvious that $\zeta = \sum_{n=0}^{\infty} \delta_{S_n}$, $S_n = \sum_{k=1}^n X_k$, does not satisfy (1).

Let I be a bounded interval. It is well-known [4, 5] that if F is transient non-arithmetic then

$$(2) \quad \lim_{t \rightarrow \infty} U(I_t) = \lim_{t \rightarrow \infty} E\zeta(I_t) = c \cdot \lambda(I),$$

where $I_t = I + t$. When $c=0$ this implies (1). Suppose $c>0$. In order to prove (1) it suffices to show

$$(3) \quad \lim_{t \rightarrow \infty} E[\zeta(I_t)^2] = c^2 \cdot \lambda(I)^2,$$

because (2) and (3) imply that

$$\lim_{t \rightarrow \infty} E[\zeta(I_t) - c\lambda(I)]^2 = 0.$$

At first we show the following relation:

$$(4) \quad E[\zeta(J)^2] = 2 \int U(J_{-s})V(J \times ds) - V(J \times J)$$

for every bounded interval J , where $V = \sum_{n=1}^{\infty} G^{n*}$. In fact

$$\begin{aligned} E[\zeta(J)^2] &= \sum_{j=0}^{\infty} \sum_{k=0}^{\infty} E[\xi_1^* \cdots \xi_j(J) \cdot \xi_1^* \cdots \xi_k(J)] \\ &= 2 \sum_{0 \leq j \leq k} E[E[\xi_1^* \cdots \xi_j(J) \cdot \xi_1^* \cdots \xi_k(J) | \xi_1, \dots, \xi_j]] \\ &\quad - \sum_{j=0}^{\infty} E[\xi_1^* \cdots \xi_j(J)^2] \\ &= 2 \sum_{0 \leq j \leq k} E[\xi_1^* \cdots \xi_j(J) \cdot \xi_1^* \cdots \xi_j^* F^{(k-j)*}(J)] \\ &\quad - \sum_{j=0}^{\infty} E[\xi_1^* \cdots \xi_j(J)^2] \\ &= 2 \sum_{0 \leq j \leq k} \int G^{j*}(J \times J_{-t})F^{(k-j)*}(dt) - \sum_{j=0}^{\infty} G^{j*}(J \times J) \\ &= 2 \int V(J \times J_{-t})U(dt) - V(J \times J). \end{aligned}$$

The assumption on G implies

$$(5) \quad \lim_{t \rightarrow \infty} V(J_t \times J_t) = 0$$

for bounded J (see [2]). On the other hand since G generates a transient random walk on R^2 we have by a theorem of Bickel and Yahav [1] that

$$(6) \quad \lim_{t \rightarrow \infty} V\{(x, y); \max(x, y) \in I_t\} = c\lambda(I).$$

Since $\lim_{t \rightarrow \infty} V(I_t \times R) = \lim_{t \rightarrow \infty} U(I_t) = c\lambda(I)$ and since V is invariant under the permutation of coordinates, it follows from (5) and (6) that

$$(7) \quad \lim_{t \rightarrow \infty} V(I_t \times (-\infty, t]) = c\lambda(I)/2.$$

Let $\{m_t\}_{t>0}$ be the family of measures defined by $m_t(A) = V(I_t \times (-A)_t)$, $A \in \mathcal{R}$. Then $\{m_t\}$ is uniformly bounded and by (5) $\lim_{t \rightarrow \infty} m_t(A) = 0$ for compact A . Thus for every bounded measurable g vanishing at $\pm\infty$ we have

$$\lim_{t \rightarrow \infty} \int g(t-s) V(I_t \times ds) = \lim_{t \rightarrow \infty} \int g(s) m_t(ds) = 0.$$

The function $h(t) = U(I_t) - c\lambda(I)1_+(t)$, where 1_+ is the indicator of $[0, \infty)$, is bounded and vanishes at infinity. Therefore it follows from (7) that

$$\begin{aligned} \lim_{t \rightarrow \infty} \int U(I_{t-s}) V(I_t \times ds) \\ &= \lim_{t \rightarrow \infty} \int h(t-s) V(I_t \times ds) + \lim_{t \rightarrow \infty} c\lambda(I) V(I_t \times (-\infty, t]) \\ &= c^2 \lambda(I)^2 / 2. \end{aligned}$$

In view of (4) and (5) this proves (3) and therefore the theorem.

References

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