

NOTE ON AN ALMOST SURE INVARIANCE PRINCIPLE FOR SOME EMPIRICAL PROCESSES

By

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1. Summary. Let $\{\xi_i\}$ be a strictly stationary sequence of random variables which are distributed uniformly over the interval $[0, 1]$ and satisfy the strong mixing (s.m.) condition

$$(1.1) \quad \alpha(n) = \sup_{A \in \mathcal{M}_{-\infty}^k, B \in \mathcal{M}_{k+n}^{\infty}} |P(A \cap B) - P(A)P(B)| \downarrow 0$$

as $n \rightarrow \infty$, where \mathcal{M}_a^b is the σ -algebra generated by ξ_a, \dots, ξ_b ($a \leq b$).

Recently, Berkes and Philipp (1977) proved an almost sure invariance principle for some empirical processes by which a functional law of the iterated logarithm for the functions of s.m. sequences, a two-dimensional functional law of the iterated logarithm, etc., are easily obtained. In this note, we shall prove that Theorem 1 in Berkes and Philipp [1] remains true under the less restrictive s.m. condition.

2. The main result. Let $F_N(s)$ ($0 \leq s \leq 1$) be the empirical distribution function defined by ξ_1, \dots, ξ_N .

Let

$$(2.1) \quad R(s, t) = [t](F_{[t]}(s) - s), \quad 0 \leq s \leq 1, \quad t \geq 0$$

where $[t]$ denotes the largest integer not exceeding t . Write

$$(2.2) \quad g_t(\alpha) = I_{[0, t]}(\alpha) - t$$

where $I_{[s, t]}(\cdot)$ denotes the indicator function of the interval $[s, t]$ and for fixed s and t with $0 \leq s < t \leq 1$, put

$$(2.3) \quad x_n(s, t) = g_t(\xi_n) - g_s(\xi_n).$$

Then, we can rewrite $R(s, t)$ as

$$(2.4) \quad R(s, t) = \sum_{j=1}^{[t]} x_j(0, s).$$

Consider the covariance function

$$(2.5) \quad \Gamma(s, t) = E g_s(\xi_1) g_t(\xi_1) + \sum_{n=2}^{\infty} E g_s(\xi_1) g_t(\xi_n) + \sum_{n=2}^{\infty} E g_t(\xi_1) g_s(\xi_n) \quad (0 \leq s, t \leq 1).$$

(It is known that under the conditions of Theorem (below) the two series in (2.5) converge absolutely for $0 \leq s, t \leq 1$.)

Let

$$(2.6) \quad \sigma^2(s, t) = \Gamma(s, s) + \Gamma(t, t) - 2\Gamma(s, t).$$

It is clear that if $\Gamma(s, t)$ is positive definite, then $\sigma^2(s, t) > 0$ for $0 \leq s < t \leq 1$. Further, let $\{K(s, t), 0 \leq s \leq 1, t \geq 0\}$ be a Kiefer process, i.e., a separable Gaussian process $K(s, t)$ on $[0, 1] \times [0, \infty)$ such that $K(0, t) = K(1, t) = K(s, 0)$ for all $0 \leq s \leq 1, t \geq 0$,

$$(2.7) \quad EK(s, t) = 0$$

and

$$(2.8) \quad EK(s, t)K(s', t') = \min(t, t')\Gamma(s, s').$$

We prove the following

Theorem. *Let $\{\xi_i\}$ be random variables defined above. Suppose that $\alpha(n) = O(n^{-3/a})$ for some a ($0 < a < 1$). Suppose that $\Gamma(s, s')$ is positive definite. Then, without changing the distribution of the empirical process $R(s, t)$ of $\{\xi_n\}$ we can redefine R on a richer probability space on which there exists a Kiefer process with covariance $\min(t, t')\Gamma(s, s')$ such that*

$$(2.9) \quad \sup_{0 \leq s \leq 1} \sup_{0 \leq t \leq T} |R(s, t) - K(s, t)| = O(T^{1/2}(\log T)^{-\lambda}) \quad a.s.$$

for some $\lambda > 0$.

3. Proof. To prove Theorem, we need some lemmas. In what follows, we denote by the letter C , with or without subscript, various absolute constants.

Lemma 1. *Let X be a random variable with finite first moment. Let $\varphi(t)$ be the characteristic function of X . Further, let Z be the standardized normal random variable. If there exist two numbers L and $T(>1)$ such that for all t ($|t| \leq T$)*

$$(3.1) \quad |\varphi(t) - e^{-t^2/2}| \leq L,$$

then for all $M(>1)$

$$(3.2) \quad \sup_u |P(X < u) - \Phi(u)| \leq C \left[M^{-1} \{E|X| + E|Z|\} + L \log MT + \frac{1}{T} \right].$$

where

$$\Phi(x) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-t^2/2} dt.$$

Proof. Let $M(>1)$ be an arbitrary number. Since $E|X| < \infty$, so for all t ($|t| \leq M^{-1}$)

$$|\varphi(t) - e^{-t^2/2}| \leq E|e^{i(X-Z)t} - 1| \leq C|t|\{E|X| + E|Z|\}.$$

Hence, we have

$$\begin{aligned} \int_{|t| \leq T} \left| \frac{\varphi(t) - e^{-t^2/2}}{t} \right| dt &= \left\{ \int_{|t| \leq M^{-1}} + \int_{M^{-1} < |t| \leq T} \right\} \left| \frac{\varphi(t) - e^{-t^2/2}}{t} \right| dt \\ &\leq C \left[\{E|X| + E|Z|\} \int_{|t| \leq M^{-1}} dt + L \int_{M^{-1} < |t| \leq T} \left| \frac{1}{t} \right| dt \right] \\ &\leq C[M^{-1}\{E|X| + E|Z|\} + L \log MT]. \end{aligned}$$

Now, (3.2) follows from Theorem 2 in [2, Chap. 5, §1], which completes the proof.

We put $l = t - s$ for any pair (s, t) ($0 \leq s < t \leq 1$).

Lemma 2. *Suppose that the conditions of Theorem are satisfied. Then there exist positive numbers γ, ρ, μ and C_0 such that $1/2 < \rho < \gamma$ and*

$$(3.4) \quad \begin{aligned} P\left(\left|\sum_{j=H+1}^{H+N} x_j(s, t)\right| \geq 3Al^{(\gamma-\rho)/2} (2N \log \log N)^{1/2}\right) \\ \leq C[\exp(-A^2 C_0^{-2} l^\rho \log \log N) + A^{-2} l^\rho N^{-1/4}] \end{aligned}$$

uniformly for all pairs (s, t) ($l \geq N^{-1/2-\mu}$) and for all $H \geq 0, A > 0$ as $N \rightarrow \infty$.

Proof. Firstly, we note that if $\alpha(n) = O(n^{-3/a})$ then we can easily find positive numbers C_0 and $\gamma' (\geq 5/9)$ such that for s, t ($0 \leq s, t \leq 1$) and for all n sufficiently large

$$(3.5) \quad \sigma_n^2(s, t) = \frac{1}{n} E \left| \sum_{j=1}^n x_j(s, t) \right|^2 \leq C_0^2 l^{\gamma'}$$

since $|x_0(s, t)| \leq 1$ and $E|x_0(s, t)| \leq Cl$.

Secondly, as $\{\xi_i\}$ is strictly stationary, we shall prove Lemma 2 in the case $H=0$.

Let

$$(3.6) \quad \gamma = \min(5/9, (7-3a)/8)$$

and choose ρ so that $1/2 < \rho < \gamma$. Let N be a sufficiently large number. Let $p = [N^{1/2}(\log N)^{-3}]$ and $k = [N/2p]$. Choose a number μ so that

$$(3.7) \quad 0 < \mu < (1-a)/2(3+a).$$

For brevity, we put

$$(3.8) \quad \chi_N = (2 \log \log N)^{1/2}$$

and $\sigma = \sigma_2(s, t) (> 0)$.

For any pair (s, t) such that $l \geq N^{-1/2-\mu}$, put

$$y_j = p^{-1/2} \sigma^{-1} \sum_{i=1}^p x_{2(j-1)p+i}(s, t) \quad (j=1, \dots, k)$$

and

$$y_{k+1}^* = p^{-1/2} \sum_{i=2kp+1}^N x_i(s, t).$$

As $\{x_j(s, t)\}$ is strictly stationary, so

$$(3.9) \quad \begin{aligned} \text{LHS of (3.4)} &\leq 2P\left(\left|\sum_{j=1}^k y_j\right| \geq Al^{(\gamma-\rho)/2} \sigma^{-1} k^{1/2} \chi_N\right) \\ &\quad + P(|y_{k+1}^*| \geq Al^{(\gamma-\rho)/2} k^{1/2} \chi_N) = 2I_1 + I_2, \quad (\text{say}). \end{aligned}$$

It follows from (3.5) that

$$(3.10) \quad \begin{aligned} I_2 &\leq A^{-2} l^{-(\gamma-\rho)} k^{-1} \chi_N^{-2} E|y_{k+1}^*|^2 \\ &\leq CA^{-2} l^{-(\gamma-\rho)} k^{-1} \chi_N^{-2} p^{-1} (N-2kp) l^{\gamma'} \leq CA^{-2} N^{-1/4} l^{\rho}. \end{aligned}$$

Now, we proceed to estimate I_1 . From (3.7), (3.9) and Lemma 1 in Yoshihara [3] we have

$$E|y_1|^4 \leq C\sigma^{-4} \{l^{4/3} + l^{1-a} p^{-1} (\log p)\} \leq C\sigma^{-4} l^{2\gamma}$$

and so from Schwartz's inequality and the fact $E|y_1|^2 = 1$ we have

$$(3.11) \quad E|y_1|^3 \leq \{E|y_1|^4\}^{1/2} \{E|y_1|^2\}^{1/2} \leq C\sigma^{-2} l^{\gamma}.$$

Hence, by Lemma 1, [2, Chap. 5, §2] and (3.11) we have that for all t ($|t| \leq (1/4)T_N$)

$$\begin{aligned} &|E\{\exp(ik^{-1/2}t \sum_{j=1}^k y_j)\} - e^{-t^2/2}| \\ &\leq |E\{\exp(ik^{-1/2}t \sum_{j=1}^k y_j)\} - \prod_{j=1}^k E\{\exp(ik^{-1/2}ty_j)\}| + |\prod_{j=1}^k E\{\exp(ik^{-1/2}ty_j)\} - e^{-t^2/2}| \\ &\leq C\{k\alpha(p) + T_N^{-1}\} \end{aligned}$$

where

$$T_N = k^{1/2} \{E|y_1|^2\}^{3/2} \{E|y_1|^3\}^{-1} \geq Ck^{1/2} \sigma^2 l^{-\gamma}.$$

Since for all N sufficiently large

$$E|k^{-1/2} \sum_{j=1}^k y_j| \leq k^{1/2} E|y_1| \leq k^{1/2} \{E|y_1|^2\}^{1/2} \leq k^{1/2}$$

so using Lemma 1 (with $M=N^2$), we have

$$\sup_z |P(k^{-1/2} \sum_{j=1}^k y_j < z) - \Phi(z)| \leq CN^{-1/4} \sigma^{-2} l^{\gamma}.$$

Hence, from the non-uniform estimate of the central limit theorem and (3.5)

$$(3.12) \quad I_1 \leq C \left[\{1 - \Phi(A l^{(\gamma-\rho)/2} \sigma^{-1} \chi_N)\} + \frac{N^{-1/4} \sigma^{-2} l^\gamma}{1 + A^2 l^{(\gamma-\rho)} \sigma^{-2}} \right] \\ \leq C [\exp(-A^2 C_0^{-2} l^\rho \log \log N) + A^{-2} N^{-1/4} l^\rho].$$

Combining (3.9), (3.10) and (3.12), we have (3.4) and the proof is completed.

The following two lemmas correspond Lemmas 5.1 and 5.2 in Berkes and Philipp [1].

Lemma 3. *If (3.4) holds, then as $k \rightarrow \infty$*

$$(3.13) \quad P(\max_{1 \leq j \leq 2^r} \sup_{s_j \leq s \leq s_{j+1}} |R(s, t_k) - R(s_j, t_k)| \geq t_k^{1/2} (\log t_k)^{-4\epsilon}) \leq C \exp(-k^{4\epsilon})$$

where $r = r_k = [\log k / \log 4]$, $t_k = [\exp(k^{1-\epsilon})]$ and $\epsilon = (\gamma - \rho) / 16$.

Proof. We write for $0 \leq s < s' \leq 1$ and integers $P (\geq 0)$, $Q (\geq 1)$

$$F(P, Q, s, s') = \left| \sum_{j=P+1}^{P+Q} x_j(s, s') \right|.$$

Put $m = [(1/2 + \mu) \log t_k / \log 2]$ and write for $s_j \leq s < s_{j+1}$

$$s = s_j + \sum_{\nu=r+1}^m \beta_\nu 2^{-\nu} + \theta 2^{-m}$$

where $\beta_\nu = 0, 1$ and $0 \leq \theta \leq 1$. We define the following events:

$$E_k(\nu, a) = \{F(0, t_k, a 2^{-\nu}, (a+1) 2^{-\nu}) \geq 2 C_0 2^{-(\gamma-\rho)\nu/2} t_k^{1/2} \chi_{t_k}\} \\ E_k = \bigcup_{r < \nu \leq m+1} \bigcup_{0 \leq a < 2^\nu} E_k(\nu, a).$$

Then, applying the same method in the proof of Lemma 5.1 in [1] and using Lemma 2 we have

$$F(0, t_k, s_j, s) \leq C t_k^{1/2} (\log t_k)^{-(\gamma-\rho)/4} \quad \text{a.s.}$$

and the proof is completed.

Lemma 4. *If (3.4) holds, then as $k \rightarrow \infty$*

$$P(\text{map}_{t_k \leq t \leq t_{k+1}} \sup_{0 \leq s \leq 1} |R(s, t) - R(s, t_k)| \geq t_k^{1/2} (\log t_k)^{-\epsilon}) \leq C k^{-2}.$$

Proof. Put $p = [(1 - \mu) \log t_k / \log 4]$ and $q = [\log(t_{k+1} - t_k) / \log 2]$. We write each integer t ($t_k \leq t \leq t_{k+1}$) in the form

$$t = t_k + \sum_{0 \leq j \leq q} \tau_j 2^j = t_k + \sum_{p < j \leq q} \tau_j 2^j + \theta t_k^{(1-\mu)/2}$$

where $\tau_j = 0, 1$ and $0 \leq \theta \leq 1$. Also, we write s ($0 \leq s \leq 1$) in the form

$$s = \sum_{\nu=1}^{\infty} \sigma_\nu 2^{-\nu} = \sum_{\nu \leq m} \sigma_\nu 2^{-\nu} + \theta 2^{-m}.$$

Further, let

$$H_k(\nu, a, j, h) = \{F(t_k + h2^{j+1}, 2^j, a^j, a2^{-\nu}, (a+1)2^{-\nu}) \geq 2C_0 2^{-(\gamma-\rho)\nu/2} (q-j)^{-2} 2^{q/2} \chi_{2q}\}$$

$$H_k = \bigcup_{p < j \leq q} \bigcup_{0 \leq \nu \leq (1+\mu)j/2} \bigcup_{0 \leq a < 2^\nu} \bigcup_{0 \leq h < 2^{q-j}} H_k(\nu, a, j, h).$$

Then, applying the same method in the proof of Lemma 5.2 in [1] and using Lemma 2, we have the lemma.

From Lemmas 3 and 4, we have the following lemma.

Lemma 5. *If the conditions of Theorem are satisfied, then*

$$(3.14) \quad \max_{t_k \leq t \leq t_{k+1}} \max_{1 \leq j \leq 2^{r_k}} \max_{s_j \leq s \leq s_{j+1}} |R(s, t) - R(s_j, t_k)| \leq C t_k^{1/2} (\log t_k)^{-2} \quad a.s.$$

where t_k and r_k are the ones defined in Lemma 3.

Proof of Theorem. Since under the conditions of Theorem the corresponding result to Proposition 3.1 in [1] is proved by the method used there, so using the Berkes and Philipp method in [1] and Lemma 5, we have the theorem.

References

- [1] Berkes, I. and Philipp, W.: *An almost sure invariance principle for the empirical distribution function of mixing random variables.* Z. Wahrscheinlichkeitstheorie verw. Gebiete **41**, 115-137 (1977).
- [2] Petrov, V. V.: *Sums of independent random variables.* Berlin. Heiderberg. Springer-Verlag (1973).
- [3] Yoshihara, K.: *Weak convergence of multidimensional empirical processes for strong mixing sequences of stochastic vectors.* Z. Wahrscheinlichkeitstheorie verw. Gebiete **33**, 133-137 (1975).

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