

LIMITING BEHAVIOR OF U -STATISTICS FOR STRONGLY MIXING SEQUENCES

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Summary. Limit theorems for some non-degenerate U -statistics are established when the underlying processes satisfy some strongly mixing condition.

1. Introduction

Let (X, \mathcal{A}) be a countably generated measurable space. Let $\{\xi_i, i \geq 0\}$ be a strictly stationary X -valued stochastic sequence defined on a probability space (Ω, \mathcal{T}, P) . Denote by F the distribution of ξ_1 and by \mathcal{M}_a^b the σ -algebra generated by ξ_a, \dots, ξ_b . The sequence is called a $*$ -mixing sequence if

$$(1.1) \quad \phi(n) = \sup_{k \geq 0} \sup_{A \in \mathcal{M}_0^k, B \in \mathcal{M}_{k+n}^\infty} \frac{|P(AB) - P(A)P(B)|}{P(A)P(B)} \longrightarrow 0 \quad (n \rightarrow \infty)$$

and it is called a ϕ -mixing sequence if

$$(1.2) \quad \phi(n) = \sup_{k \geq 0} \sup_{A \in \mathcal{M}_0^k, B \in \mathcal{M}_{k+n}^\infty} \frac{|P(AB) - P(A)P(B)|}{P(A)} \longrightarrow 0 \quad (n \rightarrow \infty).$$

Finally, the sequence is called a strongly mixing sequence if

$$(1.3) \quad \alpha(n) = \sup_{k \geq 0} \sup_{A \in \mathcal{M}_0^k, B \in \mathcal{M}_{k+n}^\infty} |P(AB) - P(A)P(B)| \longrightarrow 0 \quad (n \rightarrow \infty).$$

Next, let H be a Borel measurable kernel defined on X^m . Therefore, $H: X^m \rightarrow R$ be a measurable function such that

$$(1.4) \quad \theta(F) = \int \dots \int H(x_1, \dots, x_m) \prod_{i=1}^m dF(x_i) < \infty$$

and H is symmetric in its m arguments. A U -statistic U_n is then given by

$$(1.5) \quad U_n = \frac{1}{\binom{n}{m}} \sum_{1 \leq i_1 < \dots < i_m \leq n} H(\xi_{i_1}, \dots, \xi_{i_m}) \quad (n \geq m).$$

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A kernel H is called degenerate (for the distribution F) if for all choices of $a_i \in X (1 \leq i \leq m)$ and all $j (1 \leq j \leq m)$

$$(1.6) \quad EH(a_1, \dots, a_{j-1}, \xi_j, a_{j+1}, \dots, a_m) = 0$$

and a U -statistic is called degenerate if the corresponding kernel has this property. For every $r (r=0, 1, \dots, m)$ let

$$(1.7) \quad H_r(x_1, \dots, x_r) = \int \dots \int H(x_1, \dots, x_m) \prod_{t=r+1}^m dF(x_t)$$

so that $H_0 = \theta(F)$ and $H_m = H$. It is known that every U -statistic can be written as a finite weighted sums of degenerate ones, namely,

$$(1.8) \quad Y_n = \sum_{r=0}^m \binom{m}{r} U_n^{(r)}$$

where $U_n^{(r)}$ denotes the U -statistic obtained from the degenerate kernel

$$(1.9) \quad h_r(x_1, \dots, x_r) = \sum_{j=0}^r \binom{r}{j} (-1)^{r-j} H_j(x_1, \dots, x_j).$$

Now, let $L^2(X, F)$ be the Hilbert space of square-integrable functions with respect to F . Let $\{g_i\}$ be an orthonormal basis of $L^2(X, F)$ such that $g_0 = 1$. For each $r (1 \leq r \leq m)$ put

$$(1.10) \quad g_{i_1, \dots, i_r}(x_1, \dots, x_r) = \prod_{j=1}^r g_{i_j}(x_j).$$

It is well known that for each $r (1 \leq r \leq m)$ the system $\{g_{i_1, \dots, i_r} : 0 \leq i_1 < i_2 < \dots < i_r < \infty\}$ is a basis of the Hilbert space $L^2(X^r, F^r)$. Let $\lambda_r(i_1, \dots, i_r)$ be the Fourier coefficient of the function h_r , i.e.,

$$(1.11) \quad \lambda_r(i_1, \dots, i_r) = \int \dots \int h_r(x_1, \dots, x_r) \prod_{j=1}^r g_{i_j}(x_j) \prod_{j=1}^r dF(x_j).$$

Then, for each $r (1 \leq r \leq m)$ we have

$$(1.12) \quad h_r(x_1, \dots, x_r) = \sum_{(i_j)} \lambda_r(i_1, \dots, i_r) g_{i_1, \dots, i_r}(x_1, \dots, x_r)$$

in the L^2 -sense and by the Parseval inequality

$$(1.13) \quad \int \dots \int h_r^2(x_1, \dots, x_r) \prod_{j=1}^r dF(x_j) = \sum |\lambda_r(i_1, \dots, i_r)|^2 < \infty.$$

Put

$$(1.14) \quad \lambda(i_1, \dots, i_r, 0, \dots, 0) = \lambda_r(i_1, \dots, i_r) \quad (1 \leq r \leq m-1),$$

and

$$\lambda(i_1, \dots, i_m) = \lambda_m(i_1, \dots, i_m).$$

Let

$$(1.15) \quad \sigma_n^2 = E \left(\sum_{j=1}^n h_1(\xi_j) \right)^2$$

and

$$(1.16) \quad \sigma^2 = E h_1^2(\xi_1) + 2 \sum_{j=1}^{\infty} E h_1(\xi_1) h_1(\xi_{j+1}).$$

It is known that under conditions of Theorem 1 (below) the series in (1.16) converges absolutely and

$$(1.17) \quad \sigma_n^2 = n\sigma^2(1 + o(1)) \quad \text{as } n \rightarrow \infty.$$

Denker and Keller [2] considered the limiting behavior of $(n/m\sigma_n)(U_n - \theta(F))$ when $\{\xi_i\}$ is absolutely regular. The obtained results generalized the results in [6]. In this paper, we consider the analogous problems for some strongly mixing sequences when all of degenerate kernels $\{h_2(x_1, x_2), \dots, h_m(x_1, \dots, x_m)\}$ can be written in the form of (1.12).

2. Main results

In what follows, we use the following notations:

(i) For any random variables η we put $\|\eta\|_r = \{E|\eta|^r\}^{1/r}$ ($r \geq 1$) if $E|\eta|^r < \infty$.

(ii) $[s]$ denotes the largest integer k such that $k \leq s$.

(iii) c , with or without subscript, denotes an absolute constant which is not always the same.

Let $D[0, 1]$ be the space of functions on $[0, 1]$ that are right-continuous and have left-hand limits. We endow $D[0, 1]$ the J_1 -topology. Let $\{\xi_i\}$ be a strictly stationary X -valued sequence. Let $h_1(\cdot)$ and $\{g_{k,r}(\cdot)\}$ be the ones defined in Section 1. For positive numbers a and b we put formally

$$(2.1) \quad K_1(a) = \|h_1(\xi_1)\|_a$$

and

$$(2.2) \quad K_2(b) = \max_{2 \leq r \leq m} \sup_{k \geq 1} \|g_{k,r}(\xi_1)\|_b.$$

Define the process $\{Y_n(t) : t \geq 0\}$ by

$$(2.3) \quad Y_n(t) = \frac{n}{m\sigma_n} (U_{[nt]} - \theta(F)) \quad (t \geq 0).$$

Theorem 1. *Let $\{\xi_i\}$ be a strictly stationary X -valued sequence. Let $H: X^m \rightarrow R$ be a nondegenerate kernel. Suppose that (1.12) holds for all r ($2 \leq r \leq m$) and σ^2 , defined by (1.16), is positive. Suppose that*

$$(2.4) \quad \sum |\lambda(i_1, \dots, i_m)| < \infty$$

holds. Suppose further that one of the following sets of conditions is satisfied:

(i) $\{\xi_n\}$ is $*$ -mixing, $K_1(2) < \infty$, $K_2(4) < \infty$ and

$$(2.5) \quad \sum_{n=1}^{\infty} n^m \phi(n) < \infty.$$

(ii) $\{\xi_n\}$ is ϕ -mixing, $K_1(2) < \infty$, $K_2(2m) < \infty$ and

$$(2.6) \quad \sum_{n=1}^{\infty} n^m \phi^{1/2m}(n) < \infty.$$

(iii) $\{\xi_n\}$ is strongly mixing and there exists a $\delta (> 0)$ such that $K_1(2 + \delta/m) < \infty$, $K_1(2m + \delta) < \infty$ and

$$(2.7) \quad \sum_{n=1}^{\infty} n^m \alpha^{\delta/(2m+\delta)}(n) < \infty.$$

Then, as $n \rightarrow \infty$

$$(2.8) \quad Y_n = \{Y_n(t) : 0 \leq t \leq 1\} \xrightarrow{D} W = \{W(t) : 0 \leq t \leq 1\} \text{ in } D[0, 1]$$

where W is a Wiener process on $[0, 1]$.

Remark. Conditions (ii) and (iii) in Theorem 1 may be replaced by more flexible conditions (cf. [3], [4]).

Theorem 2. Suppose conditions of Theorem 1 are satisfied. Then, without changing its distribution we can redefine the sequence $\{\xi_n, n \geq 1\}$ on a richer probability space on which there exists a Wiener process $\{W(t) : 0 \leq t \leq 1\}$ such that with probability one

$$(2.9) \quad |t^{1/2} Y_{[t]}(1) - W(t)| = o(t^{1/2}) \quad \text{as } t \rightarrow \infty.$$

3. Auxiliary results

The following is known.

Lemma 3.1. Let $\{\zeta_i\}$ be some sequence. Let random variables X and Y be \mathcal{M}_0^* - and \mathcal{M}_{i+n}^* -measurable, respectively.

(i) If $\{\zeta_i\}$ is $*$ -mixing, $E|X| < \infty$ and $E|Y| < \infty$, then

$$(3.1) \quad |EXY - EXEY| \leq \phi(n)E|X|E|Y|,$$

which implies

$$(3.2) \quad |EXY| \leq E|XY| \leq (1 + \phi(n))E|X|E|Y|.$$

(ii) Suppose $\{\zeta_i\}$ is ϕ -mixing. Let $p (> 1)$ and $q (> 1)$ be numbers such that $p^{-1} + q^{-1} = 1$. If $E|X|^p < \infty$ and $E|Y|^q < \infty$, then

$$(3.3) \quad |EXY - EXEY| \leq 2\phi^{1/p}(n)\|X\|_p\|Y\|_q.$$

(iii) Suppose $\{\zeta_i\}$ is strongly mixing. Let $p(>1)$ and $q(>1)$ be numbers such that $\gamma = p^{-1} + q^{-1} < 1$. If $E|X|^p < \infty$ and $E|Y|^q < \infty$, then

$$(3.4) \quad |EXY - EXEY| \leq 10\alpha^{\gamma}(n) \|X\|_p \|X\|_q.$$

Next, we consider the following lemma.

Lemma 3.2. Let $\{\zeta_i\}$ be some mixing sequence of zero-mean random variables. Then, for any positive integer r

$$(3.5) \quad E \left| \sum_{1 \leq i_1 < \dots < i_r \leq n} \zeta_{i_1} \dots \zeta_{i_r} \right|^2 \leq cn^r,$$

if one of the following sets of conditions is satisfied:

(i) $\{\zeta_i\}$ is \ast -mixing, $\sup_{i \geq 1} \|\zeta_i\|_2 < \infty$ and

$$(3.6) \quad \sum n^r \phi(n) < \infty.$$

(ii) $\{\zeta_i\}$ is ϕ -mixing, $\sup_{i \geq 1} \|\zeta_i\|_{2r} < \infty$ and

$$(3.7) \quad \sum n^r \phi^{1/2r}(n) < \infty.$$

(iii) $\{\zeta_i\}$ is strongly mixing and there exists a $\delta(>0)$ for which $\sup_i \|\zeta_i\|_{2r+\delta} < \infty$ and

$$(3.8) \quad \sum n^r \alpha^{\delta/(2r+\delta)}(n) < \infty.$$

The proof of Lemma 3.2 can be carried out by using Lemma 3.1 and the method of the proof of Lemma 3 in [6] and so omitted (cf. [8], Lemma 11.4.1).

The following lemma is fundamental.

Lemma 3.3. Suppose conditions of Theorem 1 are satisfied. Then for any $t(0 < t \leq 1)$

$$(3.9) \quad E|U_{[nt]}^{(r)}|^2 \leq c(r)n^{-r} \quad (r=2, \dots, m)$$

where $c(2), \dots, c(m)$ are absolute constants which do not depend on n .

Proof. Let $r(2 \leq r \leq m)$ be arbitrary. To prove (3.9) it is enough to show that

$$(3.10) \quad E \left| \sum_{1 \leq j_1 < \dots < j_r \leq [nt]} h_r(\xi_{j_1}, \dots, \xi_{j_r}) \right|^2 \leq c(r)n^r.$$

By virtue of (1.12), (3.10) is proved if we show that

$$(3.11) \quad E \left| \sum_{1 \leq j_1 < \dots < j_r \leq [nt]} \sum_{(i)} \lambda_r(i_1, \dots, i_r) \prod_{l=1}^r g_{i_l}(\xi_{j_l}) \right|^2 \leq c(r)n^r.$$

For brevity, put $p = [nt]$,

$$(3.12) \quad \lambda_r(\{i\}) = \lambda_r(i_1, \dots, i_r)$$

and

$$(3.13) \quad G_p(\{i\}) = \sum_{1 \leq j_1 < \dots < j_r \leq p} \prod_{l=1}^r g_{i_l}(\xi_{j_l}).$$

Since for each $\{i\}$ fixed $\{g_{i_l}(\xi_j)\}$ is a mixing sequence of zero-mean random variables such that $\{g_{i_l}(\xi_j)\}$ satisfies the same mixing condition as that of $\{\xi_j\}$, by the Schwarz inequality and Lemma 3.2,

$$(3.14) \quad \begin{aligned} \text{LHS of (3.11)} &\leq \sum_{\{i\}} \sum_{\{i'\}} |\lambda(\{i\})| \cdot |\lambda(\{i'\})| \cdot E |G_p(\{i\})G_p(\{i'\})| \\ &\leq \sum_{\{i\}} \sum_{\{i'\}} |\lambda(\{i\})| \cdot |\lambda(\{i'\})| \cdot \|G_p(\{i\})\|_2 \|G_p(\{i'\})\|_2 \\ &\leq c p^r \sum_{\{i\}} |\lambda(\{i\})| \sum_{\{i'\}} |\lambda(\{i'\})|. \end{aligned}$$

Now, (3.10) follows from (3.14) and (1.13). \square

4. Proofs

Proof of Theorem 1. It is easily shown that under conditions of Theorem 1

$$(4.1) \quad \{\sigma_n^{-1} \sum_{j=1}^{[nt]} (h_1(\xi_j) - \theta(F)) : 0 \leq t \leq 1\} \xrightarrow{D} W = \{W(t) : 0 \leq t \leq 1\}$$

(cf. [3], [4]). Hence, (2.9) is obtained from (4.1) and Lemma 3.3. \square

Proof of Theorem 2. By results in [5], we can show that if condition of Theorem 1 are satisfied, then without changing its distribution we can redefine the sequence $\{\xi_n, n \geq 1\}$ on a richer probability space such that for some γ ($0 < \gamma < 1/2$)

$$(4.2) \quad |\sigma_n^{-1} \sum_{j=1}^{[t]} (h_1(\xi_j) - \theta(F)) - W(t)| = O(t^{1/2-\gamma}) \quad \text{a. s.}$$

as $t \rightarrow \infty$. On the other hand, by Lemma 3.3 we can easily show that

$$(4.3) \quad \left| \sum_{r=2}^m \binom{m}{r} U_{[t]}^{(r)} \right| = o(t^{-1/8}) \quad \text{a. s.}$$

as $t \rightarrow \infty$. Hence, (2.10) is obtained from (4.2) and (4.3). \square

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